

# Random Walks on Truncated Cubes and Sampling 0-1 Knapsack Solutions \*

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## Abstract

We solve an open problem concerning the mixing time of symmetric random walk on the  $n$ -dimensional cube truncated by a hyperplane, showing that it is polynomial in  $n$ . As a consequence, we obtain a fully-polynomial randomized approximation scheme for counting the feasible solutions of a 0-1 knapsack problem. The results extend to the case of any fixed number of hyperplanes. The key ingredient in our analysis is a combinatorial construction we call a “balanced almost uniform permutation,” which is of independent interest.

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# 1 Introduction

For a positive real vector  $\mathbf{a} = (a_i)_{i=1}^n$  and real number  $b$ , let  $\Omega$  denote the set of 0-1 vectors  $\mathbf{x} = (x_i)_{i=1}^n$  for which

$$\mathbf{a} \cdot \mathbf{x} \equiv \sum_{i=1}^n a_i x_i \leq b.$$

Geometrically, we can view  $\Omega$  as the set of vertices of the  $n$ -dimensional cube  $\{0, 1\}^n$  which lie on one side of the hyperplane  $\mathbf{a} \cdot \mathbf{x} = b$ . Combinatorially,  $\Omega$  is the set of feasible solutions to the 0-1 knapsack problem defined by  $\mathbf{a}$  and  $b$ : if we think of the  $a_i$  as the weights of a set of  $n$  items, and  $b$  as the capacity (weight limit) of a knapsack, then there is a 1-1 correspondence between vectors  $\mathbf{x} \in \Omega$  and subsets of items  $X$  whose aggregated weight does not exceed the knapsack capacity, given by  $X = \{i : x_i = 1\}$ . We shall write  $a(X)$  for the weight of  $X$ , i.e.,  $a(X) = \sum_{i \in X} a_i$ .

This paper is concerned with the natural nearest neighbor random walk on the “truncated cube”  $G_\Omega$  (i.e., the subgraph of the cube induced by  $\Omega$ ), in which the probability of moving to any neighbor is  $\frac{1}{n}$ . In the knapsack terminology, this walk is described by the following transition rule from any state  $X \in \Omega$ :

1. with probability  $\frac{1}{2}$  do nothing; else
2. pick an item  $i \in \{1, \dots, n\}$  uniformly at random (u.a.r.);
3. if  $i \in X$ , move to  $X - \{i\}$ ; if  $i \notin X$  and  $a(X \cup \{i\}) \leq b$ , move to  $X \cup \{i\}$ ; else do nothing.

(We have added a uniform holding probability of  $\frac{1}{2}$  at every state to avoid technical issues involving periodicity.)

It is easy to check that this walk converges to the uniform distribution over  $\Omega$ , for any choice of  $\mathbf{a}$  and  $b$ . Our main concern in this paper is with the number of steps required until the distribution is close to uniform, starting from an arbitrary initial state. We refer to this as the *mixing time* of the walk (see section 2 for a precise definition).

The question of determining good bounds on the mixing time of this random walk has been posed as an open problem in several places (e.g., [5, 12, 14, 17]), and is of interest for two main reasons. First, despite the development over the past decade of several powerful tools for analyzing the mixing time of combinatorial Markov chains, which have led to a string of surprising results in the field (see, e.g., [12, 9, 14, 17] for surveys), this natural and deceptively simple example remains a challenge to all existing techniques. In particular, the mixing time is not known to be bounded by any polynomial function of  $n$ . There is strong geometric intuition that it should be: random walk on the entire cube  $\{0, 1\}^n$  has mixing time only  $O(n \log n)$ , and truncation by a hyperplane presumably cannot create “bottlenecks” that would severely slow down convergence. Nonetheless, the best known bound on the mixing time remains  $\exp(O(\sqrt{n}(\log n)^{5/2}))$  [5], which beats the trivial bound of  $\exp(O(n))$  but is still exponential.

The second reason for the interest in this random walk is its connection to approximate counting. Because the walk converges to the uniform distribution over  $\Omega$ , it provides an algorithm for sampling (almost) uniformly at random from  $\Omega$ : just simulate the walk for sufficiently many steps, starting from an arbitrary vertex of  $G_\Omega$ , and output the final vertex. (This is often known as the “Markov chain Monte Carlo” paradigm.) The running time of the algorithm will be essentially the mixing time of the walk. Now by a well-known relationship based on self-reducibility, for most natural combinatorial structures (including 0-1 knapsack solutions) the problems of approximate counting and of sampling from an (almost) uniform distribution are polynomial time reducible to

one another [13, 12]. Therefore, a proof that the above random walk has mixing time polynomial in  $n$  would immediately imply the existence of a polynomial time approximation algorithm (actually a *full polynomial randomized approximation scheme*, or *fpras* [12]) for computing  $|\Omega|$ , the number of feasible knapsack solutions. This problem is  $\#P$ -complete in exact form, so an fpras is essentially the best we can hope for. Again, following a string of approximate counting algorithms for  $\#P$ -complete problems in recent years (many of them based on the Markov chain Monte Carlo technique), this problem remains one of the canonical unresolved examples.

In this paper we prove that the above random walk is indeed rapidly mixing, with a mixing time of  $O(n^{9/2+\epsilon})$  steps for any  $\epsilon > 0$ . This immediately implies the existence of the first fpras for counting 0-1 knapsack solutions.<sup>1</sup> Along the way we develop some new machinery for bounding the mixing time of Markov chains which we believe will be useful for tackling other examples of a similar flavor, and possibly beyond.

We also present a non-trivial extension of these results to the case of multiple hyperplanes (more precisely, multiple constraints of the form  $\mathbf{a}_j \cdot \mathbf{x} \leq b_j$  for non-negative vectors  $\mathbf{a}_j$ ).<sup>2</sup> Here we are also able to prove a mixing time of  $O(n^c)$  (where  $c$  is a constant) for any fixed number of hyperplanes. The exponent  $c$  depends on the number  $d$  of hyperplanes, but this is inevitable as it is not hard to prove a lower bound of  $n^{\Omega(d)}$  on the mixing time. Moreover, it is possible to encode NP-hard problems if the number of hyperplanes is permitted to depend on  $n$ , so we would not expect any polynomial time sampling algorithm for this case.

To prove rapid mixing we use a technique based on multicommodity flow (see [18]): if we can route unit flow between each pair of vertices  $X, Y$  in  $G_\Omega$  simultaneously in such a way that no edge carries too much flow, then the random walk is rapidly mixing. This technique is well known, but most previous applications (e.g., [10, 11]) have made use of “degenerate” flows in which all  $X \rightarrow Y$  flow is routed along a single canonical path (though see [2, 18] for exceptions). Our analysis seems to rely essentially on spreading out the flow along multiple paths.

The key ingredient in our analysis is the specification of these paths, which we achieve via an auxiliary combinatorial construction that we believe is of independent interest and will find further applications elsewhere.<sup>3</sup> Note that a shortest path between a pair of vertices  $X, Y$  of  $G_\Omega$  can be viewed as a *permutation* of the symmetric difference  $X \oplus Y$ , the set of items that must be added to or removed from the knapsack in passing from  $X$  to  $Y$ . A natural approach to defining a good flow is to use a *random* permutation, so that the flow is spread evenly among all shortest paths and no edge is overloaded. However, a fundamental problem with this approach is that a random permutation will tend to violate the knapsack constraint, as too many items will have been added at some intermediate point. Slightly less obviously, a symmetric problem arises because a random permutation will tend to remove too many items at some intermediate point, causing congestion among edges of the hypercube near the origin. To avoid these problems, we want our permutations to remain “balanced,” in the sense that items are added and removed at approximately the correct rates throughout the path; but we also want them to be “sufficiently random” to ensure a well spread flow. More specifically, it turns out that we require the distribution of the initial segment  $\{\pi(1), \dots, \pi(k)\}$ , viewed as an unordered set, to be “almost uniform.” We call permutations with these properties *balanced almost uniform* permutations. A main contribution of this paper is to

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<sup>1</sup>Very recently, and after the appearance of the conference version of this paper [16], Martin Dyer gave an algorithm for random sampling and approximate counting of knapsack solutions with running time only  $O(n^{2.5})$  [4]. However, Dyer’s algorithm is based on dynamic programming and gives no insight into the mixing time of the random walk.

<sup>2</sup>We mention in passing that all our results extend from the 0-1 case to more general cubes of the form  $[0, \dots, L]^n$ . This extension is purely technical and does not require any substantial new ideas, so we omit the details.

<sup>3</sup>Indeed, these ideas have already been used by Cryan *et al.* in the analysis of random walk algorithms for contingency tables [1].

show the existence of such permutations.

The remainder of the paper is structured as follows. We begin with some necessary background on flows and rapid mixing in section 2. Then in section 3 we define the notion of balanced almost uniform permutations and show how to construct them; this section is independent of the random walk analysis and should be of wider interest. We go on to use these balanced permutations to define a good flow for the knapsack random walk in section 4. The extension to multiple constraints is handled in section 5; this involves extending our construction of balanced almost uniform permutations in a non-trivial way from scalar weights to vectors in arbitrary dimension, which is again of independent interest. We conclude with the proofs of some technical lemmas in section 6.

## 2 The mixing time and multicommodity flow

As indicated earlier, we will view elements of  $\Omega$  either as 0-1 vectors  $\mathbf{x} = (x_i)_{i=1}^n$  or, more commonly, as subsets  $X \subseteq \{1, \dots, n\}$ , under the equivalence  $X = \{i : x_i = 1\}$ . Recall that  $a(X) = \sum_{i \in X} a_i$  is the weight of  $X$ , so that  $\Omega = \{X : a(X) \leq b\}$ . Without loss of generality, we will assume that  $a_i \leq b$  for all  $i$ .

We consider the symmetric random walk on the portion  $G_\Omega$  of the hypercube  $\{0, 1\}^n$  defined in the Introduction. This walk is connected (all states communicate via the zero vector) and aperiodic (because of the holding probabilities), and since the transition probabilities are symmetric, the distribution at time  $t$  converges to the uniform distribution over  $\Omega$  as  $t \rightarrow \infty$ , regardless of the initial state. Our goal is to bound the rate of convergence as measured by the *mixing time*, defined as

$$\tau_{\text{mix}} = \max_{X_0} \min \left\{ t : \|P_t - \mathcal{U}\| \leq \frac{1}{4} \right\},$$

where  $X_0$  is the initial state,  $P_t$  is the distribution of the walk at time  $t$ ,  $\mathcal{U}$  is the uniform distribution over  $\Omega$ , and  $\|\cdot\|$  denotes variation distance.<sup>4</sup> Thus  $\tau_{\text{mix}}$  is the number of steps required, starting from any initial state, to get the variation distance from the uniform distribution down to  $\frac{1}{4}$ . By standard facts about geometric convergence,  $O(\tau_{\text{mix}} \log \epsilon^{-1})$  steps suffice to reduce the variation distance to any desired  $\epsilon$ .

Fairly standard techniques (see [18]) allow us to estimate  $\tau_{\text{mix}}$  by setting up a suitable multi-commodity flow on the underlying graph  $G_\Omega$ . Our task is to route one unit of flow from  $X$  to  $Y$ , for each ordered pair of vertices  $X, Y \in \Omega$  simultaneously. For any such flow  $f$ , and any oriented edge  $e$  in  $G_\Omega$ , let  $f(e)$  denote the total flow along  $e$ ; i.e.,  $f(e)$  is the sum over all ordered pairs  $X, Y$  of the  $X \rightarrow Y$  flow carried by  $e$ . Define the *congestion*  $\mathcal{C}(f) = \frac{1}{|\Omega|} \max_e f(e)$ , i.e., the maximum flow along any edge normalized by  $|\Omega|$ , and the *length*  $\mathcal{L}(f)$  to be the length of a longest flow-carrying path. The following theorem<sup>5</sup> is a special case of results in [18], which in turn generalize those in [3]:

**Theorem 2.1** *For any flow  $f$ , the mixing time is bounded by  $\tau_{\text{mix}} \leq 4n(n+1)\mathcal{C}(f)\mathcal{L}(f)$ .*

We will bound  $\tau_{\text{mix}}$  by constructing a flow  $f$  with congestion  $\mathcal{C}(f) = O(n^{3/2+\epsilon})$  for any  $\epsilon > 0$ , and length  $\mathcal{L}(f) = O(n)$ . By Theorem 2.1 this implies  $\tau_{\text{mix}} = O(n^{9/2+\epsilon})$ .

<sup>4</sup>For probability distributions  $\mu, \nu$  on  $\Omega$ , the variation distance is defined as  $\|\mu - \nu\| = \frac{1}{2} \sum_{x \in \Omega} |\mu(x) - \nu(x)| = \max_{S \subseteq \Omega} |\mu(S) - \nu(S)|$ .

<sup>5</sup>This theorem applies to symmetric random walk on *any* connected subgraph of the hypercube  $\{0, 1\}^n$ , in which transitions are made to each neighbor with probability  $\frac{1}{2n}$ .

**Remark:** We note that our bound on the mixing time is only slightly larger than the upper bound of  $O(n^3)$  which one obtains by applying Theorem 2.1 to the hypercube itself (without the hyperplane constraint): see, e.g., [19]. This is in turn somewhat off from the true mixing time of  $O(n \log n)$ . On the other hand, it is fairly easy to obtain a lower bound of  $\Omega(n^2/\log n)$  for the mixing time of the truncated cube: consider, for example, an instance in which  $\log n$  items have weight 1, the other  $n - \log n$  items have weight  $n$ , and the knapsack capacity is  $b = n$ .  $\square$

As explained in the Introduction, our flow will be based on the idea of a balanced almost uniform permutation. We devote the next section to this topic and then return to the knapsack random walk in section 4.

### 3 Balanced almost uniform permutations

We begin by defining the notions of “balanced” and “almost uniform” permutations. We will write  $\mathcal{S}_m$  to denote the set of all permutations of  $\{1, \dots, m\}$ .

**Definition 3.1** *Let  $\{w_i\}_{i=1}^m$  be a set of real (not necessarily positive) weights, with  $M = \max_{i \leq m} |w_i|$  and  $W = \sum_i w_i$ , and let  $\Delta \geq 1$  be a nonnegative number. A permutation  $\pi \in \mathcal{S}_m$  is  $\Delta$ -balanced if, for all  $k$  with  $1 \leq k \leq m$ ,*

$$\min\{W, 0\} - \Delta M \leq \sum_{i=1}^k w_{\pi(i)} \leq \max\{W, 0\} + \Delta M. \quad (1)$$

Thus a balanced permutation is one whose partial sums do not fluctuate widely. In particular, if  $\sum_i w_i = 0$  then condition (1) becomes  $|\sum_{i=1}^k w_{\pi(i)}| \leq \Delta M$ .

**Definition 3.2** *Let  $\pi$  be a random permutation in  $\mathcal{S}_m$ , and let  $\lambda \in \mathbf{R}$ . We call  $\pi$  a  $\lambda$ -uniform permutation if*

$$\Pr[\pi\{1, \dots, k\} = U] \leq \lambda \times \binom{m}{k}^{-1} \quad (2)$$

*for every  $k$  with  $1 \leq k \leq m$  and every  $U \subseteq \{1, \dots, m\}$  of cardinality  $k$ . (Here  $\pi\{1, \dots, k\}$  denotes the initial segment  $\{\pi(1), \dots, \pi(k)\}$ .)*

Note that, if  $\pi$  were a uniform random permutation, the probability in (2) would be exactly  $\binom{m}{k}^{-1}$  for every  $U$ . In a  $\lambda$ -uniform permutation the probabilities are permitted to vary with  $U$ , but only by an amount specified by the parameter  $\lambda$ . In our applications,  $\lambda$  will be a fixed polynomial function of  $m$ ; in this case we call  $\pi$  an *almost uniform* permutation. The perhaps surprising result of this section is that, for any set of weights  $\{w_i\}$ , it is possible to construct an almost uniform permutation that is guaranteed to be balanced.

#### 3.1 The bounded ratio case

In this section, we prove that there are balanced almost uniform permutations when the ratios of the weights are bounded by a constant. In section 3.2 we will show how to dispense with any restrictions on the weights.

**Lemma 3.3** *Let  $\{w_i\}_{i=1}^m$  be an arbitrary set of weights with  $1 \leq |w_i| \leq B$ , for a constant  $B \in [1, 2]$ . Then there exists a 1-balanced permutation  $\pi$  on  $\{w_i\}$  which is  $p(m)$ -uniform, where  $p(m) = Cm^{10(B-1)^2+1/2}$  for a universal constant  $C$ .*

**Proof:** Let  $M = \max_i |w_i|$  and  $W = \sum_{i=1}^m w_i$ . Assume first that  $W = 0$ ; we will show how to discharge this assumption later. Let  $I_1 = \{i : w_i > 0\}$ ,  $I_2 = \{i : w_i < 0\}$ ,  $m_1 = |I_1|$  and  $m_2 = |I_2|$ . Define the means  $\mu_1 = \frac{1}{m_1} \sum_{i \in I_1} w_i$  and  $\mu_2 = -\frac{1}{m_2} \sum_{i \in I_2} w_i$ . Note that  $m_1 \mu_1 = m_2 \mu_2$  since  $W = 0$ .

Consider an arbitrary permutation  $\nu \in \mathcal{S}_m$ . This induces permutations  $\nu_1, \nu_2$  on  $I_1, I_2$  respectively.<sup>6</sup> We call  $\nu_1$   $\alpha$ -good if, for every  $k_1$  with  $1 \leq k_1 \leq m_1$ ,

$$\left| \sum_{i=1}^{k_1} w_{\nu_1(i)} - k_1 \mu_1 \right| \leq \alpha(M-1) \sqrt{k_1^*}, \quad (3)$$

where  $k_1^* = \min\{k_1, m_1 - k_1\}$ , with an analogous definition for  $\nu_2$ . We call  $\nu$   $\alpha$ -good if both  $\nu_1$  and  $\nu_2$  are  $\alpha$ -good. Thus in a good permutation, the partial sums of both positive and negative weights are reasonably close to their expected values.

Now suppose  $\nu$  is chosen u.a.r. from  $\mathcal{S}_m$ . A routine application of Hoeffding's bound to the partial sums (see Lemma 6.1 in section 6) yields

$$\Pr[\nu \text{ is not } \alpha\text{-good}] \leq 2m \exp(-2\alpha^2). \quad (4)$$

If we set  $\alpha = \sqrt{\ln m}$ , this probability is at most  $\frac{2}{m} \leq \frac{1}{2}$  for  $m \geq 4$ .

Consider now a modified sample space in which  $\nu$  is selected u.a.r. among all  $\sqrt{\ln m}$ -good permutations. We shall write  $\Pr_{\text{unif}}$  for probabilities in the original uniform space to distinguish them from those in this modified space. By the above calculation, for any event  $\mathcal{E} \subseteq \mathcal{S}_m$  we have

$$\Pr[\mathcal{E}] \leq 2 \Pr_{\text{unif}}[\mathcal{E}]. \quad (5)$$

We are now in a position to construct our balanced almost uniform permutation. Let  $\nu$  be chosen u.a.r. from all  $\sqrt{\ln m}$ -good permutations, and let  $\nu_1, \nu_2$  be the induced permutations on  $I_1, I_2$ . To get a balanced permutation  $\pi$ , we *interleave*  $\nu_1$  and  $\nu_2$  as follows. We take the first element from  $\nu_1$ , i.e., set  $\pi(1) = \nu_1(1)$ . Thereafter, for each  $k > 1$  in turn we set  $\pi(k)$  to be the next element in  $\nu_2$  if  $\sum_{i=1}^{k-1} w_{\pi(i)} > 0$ , and the next element in  $\nu_1$  otherwise. Since  $\sum_i w_i = 0$  this process is well-defined and yields a permutation  $\pi \in \mathcal{S}_m$ . Moreover, since  $|w_i| \leq M$  for all  $i$  it is clear that  $\pi$  satisfies the balance condition (1) with  $\Delta = 1$ .

We now need to verify the uniformity condition (2), for  $\lambda = p(m)$ . Let  $U \subseteq \{1, \dots, m\}$  be any subset of items with  $|U| = k$  and  $\Pr[\pi\{1, \dots, k\} = U] > 0$  and define

$$U_1 = U \cap I_1; \quad U_2 = U \cap I_2; \quad k_1 = |U_1|; \quad k_2 = |U_2|.$$

Then we have

$$\begin{aligned} \Pr[\pi\{1, \dots, k\} = U] &\leq \Pr[\nu_1\{1, \dots, k_1\} = U_1 \text{ and } \nu_2\{1, \dots, k_2\} = U_2] \\ &\leq 2 \Pr_{\text{unif}}[\nu_1\{1, \dots, k_1\} = U_1 \text{ and } \nu_2\{1, \dots, k_2\} = U_2] \\ &= \frac{2}{\binom{m_1}{k_1} \binom{m_2}{k_2}}, \end{aligned} \quad (6)$$

where the second inequality follows from (5).

Now some routine calculations involving Stirling's formula (see Lemma 6.2 in section 6) allow us to relate  $\binom{m_1}{k_1} \binom{m_2}{k_2}$  to  $\binom{m_1+m_2}{k_1+k_2} = \binom{m}{k}$ . Specifically, (6) becomes

$$\Pr[\pi\{1, \dots, k\} = U] \leq \frac{Am^{1/2}}{\binom{m}{k}} \exp \left\{ \frac{l^2 + \frac{1}{2}|l|}{\gamma(1-\gamma)} \left( \frac{1}{m_1} + \frac{1}{m_2} \right) \right\}, \quad (7)$$

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<sup>6</sup>Formally, we view  $\nu_1$  as a bijection from  $\{1, \dots, m_1\}$  to  $I_1$ , and similarly for  $\nu_2$ . Throughout we shall adopt this convention where appropriate, without comment.

where  $\gamma = \frac{k}{m}$ ,  $l = \frac{m_1 k_2 - m_2 k_1}{m}$ , and  $A > 0$  is a universal constant. The quantity  $l$  measures the deviation of the numbers  $k_1, k_2$  of positive and negative elements in  $U$  from the “expected” values  $\gamma m_1, \gamma m_2$  respectively. But since  $\pi$  is balanced,  $\nu$  is good, and the element sizes do not vary too much,  $|l|$  cannot in fact be very large. To formalize this intuition, note first that

$$l = (k_2 \mu_2 - k_1 \mu_1) \frac{m_2}{\mu_1 m}, \quad (8)$$

since  $\frac{m_2}{m_1} = \frac{\mu_1}{\mu_2}$ . Now by the goodness condition (3) on  $\nu_1, \nu_2$  we have

$$\begin{aligned} \left| \sum_{i=1}^k w_{\pi(i)} - (k_1 \mu_1 - k_2 \mu_2) \right| &= \left| \left( \sum_{i=1}^{k_1} w_{\nu_1(i)} + \sum_{i=1}^{k_2} w_{\nu_2(i)} \right) - (k_1 \mu_1 - k_2 \mu_2) \right| \\ &\leq 2(M-1) \sqrt{k^* \ln m}, \end{aligned}$$

where  $k^* = \min\{k, m-k\}$ . Since  $\pi$  is 1-balanced we also know that  $|\sum_{i=1}^k w_{\pi(i)}| \leq M$ , and therefore

$$|k_1 \mu_1 - k_2 \mu_2| \leq 2(M-1) \sqrt{k^* \ln m} + M.$$

Together with (8) and our assumption that  $M \leq B$ , this implies the following bound on  $|l|$ :

$$|l| \leq \left( 2(B-1) \sqrt{k^* \ln m} + B \right) \frac{m_2}{\mu_1 m}.$$

Plugging in this value for  $|l|$ , the exponent in (7) is bounded above, when  $|l|$  is sufficiently large, by

$$\begin{aligned} &5(B-1)^2 k^* \ln m \frac{m_2^2}{\mu_1^2 m^2} \frac{m^2}{k(m-k)} \frac{m}{m_1 m_2} \\ &= 5(B-1)^2 \ln m \frac{k^* m}{k(m-k)} \frac{1}{\mu_1 \mu_2} \\ &\leq 10(B-1)^2 \ln m, \end{aligned} \quad (9)$$

since  $k(m-k) \geq \frac{k^* m}{2}$  and  $\mu_1, \mu_2 \geq 1$ . Note also that if  $|l|$  is bounded, then so is the exponent in (7), since

$$\frac{1}{\gamma(1-\gamma)} \left( \frac{1}{m_1} + \frac{1}{m_2} \right) = \frac{m^2}{k(m-k)} \frac{m}{m_1 m_2} \leq \frac{m^3}{(m/2)(m/3)(m/3)} \leq 18,$$

where we have used the fact that  $m_1, m_2 \geq \frac{m}{3}$  which follows because  $B \leq 2$  and  $W = 0$ .

Thus (7) becomes

$$\Pr[\pi\{1, \dots, k\} = U] \leq C \binom{m}{k}^{-1} m^{10(B-1)^2 + 1/2}, \quad (10)$$

for a universal constant  $C$ , which verifies the uniformity condition (2) with  $\lambda = p(m) = C m^{10(B-1)^2 + 1/2}$ .

This concludes the proof of the theorem for the case  $W = \sum_i w_i = 0$ . We can extend the argument to general values of  $W$  using a simple trick. We will assume  $W > 0$ ; the case  $W < 0$  is entirely symmetrical. We begin by padding the sequence of weights with  $d = \lceil W/M \rceil$  values  $w_{m+1}, \dots, w_{m+d}$  each of which (except possibly the last) is  $-M$ , so that  $\sum_{i=1}^{m+d} w_i = 0$ . Note that  $d \leq m$ . By the above argument for the  $W = 0$  case, we can construct a 1-balanced almost uniform permutation  $\pi'$  on this padded sequence (though see the remark immediately following this proof). Let  $\pi$  be the *induced* permutation on the weights  $\{w_i\}_{i=1}^n$ . We claim that  $\pi$  is also 1-balanced and almost uniform.

To see that  $\pi$  is 1-balanced, note that

$$\begin{aligned} \sum_{i=1}^k w_{\pi(i)} &\geq \sum_{i=1}^{k'} w_{\pi'(i)} \geq -M; \quad \text{and} \\ \sum_{i=1}^k w_{\pi(i)} &\leq \sum_{i=1}^{k'} w_{\pi'(i)} + W \leq M + W, \end{aligned}$$

for some  $k' \geq k$ , using the balance property of  $\pi'$ .

To see that  $\pi$  is almost uniform, let us call the indices  $\{1, \dots, m\}$  *true* and the remainder *fake*. Let  $U$  be an arbitrary subset of true indices of cardinality  $k$ . We need to show that

$$\Pr[\pi\{1, \dots, k\} = U] \leq \binom{m}{k}^{-1} p(m). \quad (11)$$

Since  $\pi$  is induced by  $\pi'$ , this probability is bounded above by  $\sum_S \Pr[\mathcal{E}_S]$ , where for  $S \subseteq \{1, \dots, m+d\}$ ,  $\mathcal{E}_S$  is the event that  $\pi'\{1, \dots, |S|\} = S$  and the sum is over all  $S$  of the form  $U \cup U'$ , where all elements of  $U'$  are fake. Now by the almost uniformity of  $\pi'$ , this sum is at most

$$p(m+d) \sum_S \Pr_{\text{unif}}[\mathcal{E}_S], \quad (12)$$

where  $\Pr_{\text{unif}}$  denotes probability under the uniform distribution on permutations in  $\mathcal{S}_{m+d}$ . But the sum in (12) is just the expectation, under the uniform distribution, of the random variable  $X = \sum_S X_S$ , where  $X_S$  is the indicator r.v. of  $\mathcal{E}_S$ . Thus  $X$  counts the number of events  $\mathcal{E}_S$  that occur. We claim that

$$\mathbb{E}(X) = \binom{m}{k}^{-1} \left(1 + \frac{d}{m+1}\right). \quad (13)$$

This will complete the verification of condition (11); for replacing the sum in (12) by  $\mathbb{E}(X)$  gives

$$\begin{aligned} \Pr[\pi\{1, \dots, k\} = U] &\leq \binom{m}{k}^{-1} \left(1 + \frac{d}{m+1}\right) p(m+d) \\ &\leq \binom{m}{k}^{-1} 2C(2m)^{10(B-1)^2+1/2}, \end{aligned}$$

where the second inequality holds because  $d \leq m$ . Thus, if we incorporate into  $C$  an extra factor of  $2^{10(2-1)^2+3/2}$ , we see that  $\pi$  is  $\lambda$ -balanced for  $\lambda = p(m) = Cm^{10(B-1)^2+1/2}$ , as required.

To see the claim in (13), let  $\mathcal{E}$  be the event that  $\pi\{1, \dots, k\} = U$ . Clearly  $\Pr_{\text{unif}}[\mathcal{E}] = \binom{m}{k}^{-1}$ , and  $X = 0$  unless  $\mathcal{E}$  occurs, so we have

$$\mathbb{E}(X) = \binom{m}{k}^{-1} \mathbb{E}(X|\mathcal{E}). \quad (14)$$

Conditioning now on  $\mathcal{E}$ , let  $r$  be the position in  $\pi'$  of the last element of  $U$ , so that  $U \subseteq \pi'\{1, \dots, r\}$  and  $\pi'(r) \in U$ . Also, let  $s$  be the position of the *next* true element, i.e.,  $\pi'(s)$  is true and  $\pi'(t)$  is fake for  $r < t < s$ . (If no such element exists, let  $s = m+d+1$ .) Then  $\mathcal{E}_S$  holds for precisely those sets  $S = \pi'\{1, \dots, t\}$ , where  $r \leq t < s$ . The number of such sets is just the number of fake elements that fall between the true element at position  $r$  and the next true element (at position  $s$ ), plus one. The expectation of this quantity under the uniform distribution is plainly  $1 + \frac{d}{m+1}$ . Plugging this into (14) we get the value claimed in (13), which concludes the proof that  $\pi$  is almost uniform.

□

**Remark:** We should point out that the padded sequence we introduced in the second part of the above proof might contain one weight whose absolute value is less than one. Thus it is not, in a strict sense, a special case of the earlier  $W = 0$  case, where we assumed that all the weights had absolute values in the range  $[1, B]$ . However, as the reader may easily verify, the analysis leading up to equation (10) still holds (with minor modifications) even when there is a single small weight.  $\square$

### 3.2 The general case

In this section we extend our construction of balanced almost uniform permutations to handle arbitrary weights. The chief obstacle here is that it is no longer true (as in the bounded ratio case) that each item of positive weight can be balanced by a bounded number of items of negative weight. To overcome this difficulty, we will need to group items into “intervals” so that each interval has approximately the same (positive or negative) weight. We can then reduce to the bounded ratio case.

The following theorem is a generalization of Lemma 3.3; it says that we can construct a balanced almost uniform permutation for an arbitrary set of weights. Moreover, we can bound the uniformity parameter  $\lambda$  by a polynomial whose degree is arbitrarily close to  $1/2$  at the cost of a modest increase in the balance parameter  $\Delta$ . This is almost the best that one can hope for: we encourage the reader to check by a simple counting argument that, if we have  $m/2$  weights of  $+1$  and  $m/2$  of  $-1$ , then for any constants  $\Delta, C$  and  $p < 1/2$ , there can be no  $\Delta$ -balanced  $Cm^p$ -uniform permutation if  $m$  is sufficiently large.

In order to achieve the tightest possible bound in our random walk analysis in the next subsection, we shall actually prove a slightly stronger uniformity property which can be obtained with no additional effort. Call  $\pi$  *strongly  $\lambda$ -uniform* if

$$\Pr[\pi\{1, \dots, k\} = U \text{ and } \pi(k+1) = l] \leq \lambda \times \binom{m}{k \ m-k-1 \ 1}^{-1} \quad (15)$$

for every  $k$  with  $1 \leq k \leq m$ , every  $U \subseteq \{1, \dots, m\}$  of cardinality  $k$ , and every  $l \notin U$ . Note that the expression on the right-hand side of (15) is just  $\lambda$  times the probability of the given event if  $\pi$  were chosen uniformly at random. Plainly (15) is a strengthening of equation (2) in Definition 3.2.

**Theorem 3.4** *Fix  $0 < \epsilon < 2$  and let  $\Delta = 1 + \sqrt{90/\epsilon}$ . For any  $m$  and set of weights  $\{w_i\}_{i=1}^m$ , there exists a  $\Delta$ -balanced strongly  $Cm^{1/2+\epsilon}$ -uniform permutation, where  $C$  is a universal constant.*

To illustrate this theorem, if we plug in the value  $\epsilon = 3/2$  (say) we obtain the following corollary:

**Corollary 3.5** *For any set of weights  $\{w_i\}_{i=1}^m$ , there exists a  $Cm^2$ -uniform, 9-balanced permutation.*

**Proof of Theorem 3.4:** First, we observe that the permutation constructed in the proof of Lemma 3.3 actually satisfies the strong uniformity property (15), provided that we incorporate an extra factor of 3 into the constant  $C$ . To see why, note that the permutation must first choose  $U$  and then  $l$ . Suppose that the sum of the weights of the indices of  $U$  is positive. Then  $l$  is chosen uniformly from the remaining indices of negative weight. But these must constitute at least  $1/3$  of all of the remaining indices since  $B \leq 2$ . Hence the restriction on the choice of  $l$  introduces a factor of at most 3 into the uniformity parameter.

Now let  $\{w_i\}$  be an arbitrary set of weights, let  $M = \max_i |w_i|$  and set  $\widehat{\Delta} = \frac{\Delta-1}{3}$ . Let  $\beta$  be a uniform random permutation in  $\mathcal{S}_m$ , and  $T_1$  be the smallest  $t$  such that the partial sum  $\sum_{i=1}^t w_{\beta(i)}$  has absolute value greater than  $\widehat{\Delta}M$  (or  $T_1 = m$  if no such  $t$  exists). Similarly, let  $T_2$  be the smallest  $t > T_1$  such that  $|\sum_{i=T_1+1}^t w_{\beta(i)}| > \widehat{\Delta}M$ . Define  $T_3, T_4, \dots$  in the same way. Then let  $I_1$  be the sequence  $\{\beta(i)\}_{i=1}^{T_1}$ , and  $I_2$  the sequence  $\{\beta(T_1 + i)\}_{i=1}^{T_2-T_1}$ . Continue in this way, dividing  $\beta$  into intervals  $I_1, \dots, I_q$  (so that  $T_q = m$ ).

Now let  $\alpha_i$  be the aggregated weight of interval  $I_i$  for  $i = 1, 2, \dots, q-1$ . Note that  $|\alpha_i| \in [\widehat{\Delta}M, (\widehat{\Delta} + 1)M]$  for all  $i < q$ , so the ratio of the weights of any two of these intervals is at most  $(\widehat{\Delta} + 1)/\widehat{\Delta} \in [1, 2]$ . Thus, by Lemma 3.3, there exists a 1-balanced  $\lambda$ -uniform permutation on  $\{\alpha_i\}_{i=1}^{q-1}$  for  $\lambda = Cq^{10}((\widehat{\Delta}+1)/\widehat{\Delta}-1)^{2+1/2} = Cq^{1/2+\epsilon}$ , where  $C$  is a universal constant. Moreover, as argued above, we can in fact assume that this permutation is *strongly*  $\lambda$ -uniform. Call this permutation  $\pi_I$ . We claim that the permutation

$$\pi = I_{\pi_I(1)} I_{\pi_I(2)} \cdots I_{\pi_I(q-1)} I_q$$

obtained by rearranging the first  $q-1$  intervals according to  $\pi_I$  is a  $\Delta$ -balanced strongly  $Cm^{1/2+\epsilon}$ -uniform permutation on the original  $m$  weights.

We prove the balance property first. Let  $W' = \sum_{i=1}^{q-1} \alpha_i = W - \alpha_q$ . Since  $\pi_I$  is 1-balanced,  $\pi$  satisfies

$$\min\{0, W'\} - (\widehat{\Delta} + 1)M \leq \sum_{i=1}^{T_j} w_{\pi(i)} \leq \max\{0, W'\} + (\widehat{\Delta} + 1)M$$

for all  $1 \leq j \leq q$ . Hence we have, for all  $j$ ,

$$\min\{0, W'\} - (2\widehat{\Delta} + 1)M \leq \sum_{i=1}^j w_{\pi(i)} \leq \max\{0, W'\} + (2\widehat{\Delta} + 1)M,$$

since the partial sums within any interval lie in the range  $[-\widehat{\Delta}M, \widehat{\Delta}M]$ . Finally, note that  $|W - W'| = |\alpha_q| \leq \widehat{\Delta}M$ . It follows that for all  $j$ ,

$$\min\{0, W\} - (3\widehat{\Delta} + 1)M \leq \sum_{i=1}^j w_{\pi(i)} \leq \max\{0, W\} + (3\widehat{\Delta} + 1)M,$$

and hence  $\pi$  is  $\Delta$ -balanced.

To verify the strong uniformity property, consider first an alternative experiment in which the permutation  $\pi_I$  is chosen u.a.r. from  $\mathcal{S}_{q-1}$ , without regard to the balance property. Note that, conditional on the value of  $q$ , the distribution of  $(I_1, \dots, I_{q-1})$  is exchangeable. Thus, re-arranging the intervals according to a uniform  $\pi_I$  is a *measure-preserving* transformation, so  $\pi$  itself has the uniform distribution. Thus we need to show that for any  $U$  and any index  $l \notin U$ , the likelihood ratio

$$\frac{\Pr[\pi\{1, \dots, k\} = U \text{ and } \pi(k+1) = l]}{\Pr_{\text{unif}}[\pi\{1, \dots, k\} = U \text{ and } \pi(k+1) = l]} \leq Cm^{1/2+\epsilon},$$

where we write  $\Pr_{\text{unif}}$  for the probability when  $\pi_I$  is uniform and  $\Pr$  for the probability when  $\pi_I$  is  $Cm^{1/2+\epsilon}$ -uniform. In fact, it suffices to show that the above bound on the likelihood ratio holds conditional on any  $\beta$ . So fix a permutation  $\beta$ . In order for the numerator to be non-zero, only the interval containing  $l$  can contain elements from both  $U$  and  $U^c$  (the complement of  $U$ ). Additionally, in the interval containing  $l$ , all the elements before  $l$  must be in  $U$  and all those

after  $l$  must be in  $U^c$ . Let  $A_1$  be the collection of intervals in  $\{I_i\}_{i=1}^{q-1}$  containing only elements of  $U$ , and  $A_2$  the collection of intervals containing only elements of  $U^c$ . Then  $|A_1| + |A_2|$  must have value either  $q - 1$  or  $q - 2$ . Writing  $\mathcal{E}_1$  for the event  $\pi_I\{1, \dots, |A_1|\} = A_1$  and  $\mathcal{E}_2$  for the event  $\pi_I\{q - 1, \dots, q - |A_2|\} = A_2$ , the above likelihood ratio is

$$\frac{\Pr[\mathcal{E}_1 \text{ and } \mathcal{E}_2]}{\Pr_{\text{unif}}[\mathcal{E}_1 \text{ and } \mathcal{E}_2]} \leq Cq^{1/2+\epsilon} \leq Cm^{1/2+\epsilon}.$$

In the case where  $|A_1| + |A_2| = q - 1$  this is just the  $Cq^{1/2+\epsilon}$ -uniformity property; when  $|A_1| + |A_2| = q - 2$  it is the strong  $Cq^{1/2+\epsilon}$ -uniformity property. Thus  $\pi$  is strongly  $Cm^{1/2+\epsilon}$ -uniform, and the proof is complete.  $\square$

## 4 Constructing a good flow

We now return to the random walk for the knapsack problem, and show how to construct a flow  $f$  with small congestion  $\mathcal{C}(f)$  and length  $\mathcal{L}(f)$ . Our construction will make heavy use of the balanced almost uniform permutations discussed in the previous section. In the first subsection below, we present a slightly informal, high-level sketch of the argument, concentrating on showing how the balance and almost uniformity properties of the permutations are used. In section 4.2 we will give the full details of the flow.

### 4.1 High-level sketch

Let  $X, Y$  be two arbitrary vertices of  $G_\Omega$ , viewed as subsets of  $\{1, \dots, n\}$ . We need to specify how to route one unit of flow from  $X$  to  $Y$ . The high-level idea is the following. Let  $S = X \oplus Y$  (where  $\oplus$  denotes symmetric difference) and  $m = |S|$ , and let  $\{w_i\}_{i=1}^m$  be an arbitrary enumeration of the weights of the items in  $S$ , with the weights of items in  $X, Y$  appearing with negative and positive signs respectively. Note that each permutation of these items specifies a geodesic path from  $X$  to  $Y$ . Thus a balanced almost uniform permutation on the  $\{w_i\}$  specifies a scheme for routing unit flow between  $X$  and  $Y$  along geodesic paths: the amount of flow allocated to a given path is just the probability assigned to the corresponding permutation. Moreover, the weight of the  $k$ th point along the path is given by  $a(X) + \sum_{i=1}^k w_{\pi(i)}$ . We would like to use the balance property to argue that these paths remain within  $\Omega$ , and the almost uniformity property to argue that they are “well spread out” and thus lead to small congestion. In what follows, we assume the existence of  $\Delta$ -balanced,  $p(m)$ -uniform permutations on any set of weights  $\{w_i\}_{i=1}^m$ , where  $\Delta$  is a fixed positive integer and  $p$  a fixed polynomial. (By Corollary 3.5 we know we can take, e.g.,  $\Delta = 9$  and  $p(m) = O(m^2)$ .)

The problem with the above idea is that we need  $X$  and  $Y$  to be at least some small distance below the bounding hyperplane, to accommodate the fluctuations that are still present within the balanced permutations. At this point, for the purposes of illustration, we make a simplifying assumption: we assume that there exists a fixed constant  $h$  (independent of  $n$  and the  $a_i$ ) such that, by removing at most  $h$  items from  $X \oplus Y$ , we arrive at vertices  $X'$  and  $Y'$  which satisfy  $a(X'), a(Y') \leq b - \Delta M$ , where  $M = \max_{i \in X' \oplus Y'} a_i$ . Note that removing the  $h = 2\Delta$  heaviest items from  $X \oplus Y$  ensures this for at least one of  $X', Y'$ . In fact, if we increase  $h$  by a constant factor then the same holds for both  $X'$  and  $Y'$  for most pairs  $(X, Y) \in \Omega \times \Omega$ . For now, however, we simply assume this property for all pairs. Denote by  $X_0, Y_0$  the sets of items removed from  $X, Y$  respectively (i.e.,  $X = X' \cup X_0$  and  $Y = Y' \cup Y_0$ , with  $|X_0 \cup Y_0| \leq h$ ).

We can now describe the flow from  $X$  to  $Y$  in three stages:

Stage 1: Send the entire unit flow along a single path from  $X$  to  $X'$  by removing the items in  $X_0$  in index order.

Stage 2: Distribute the unit flow along geodesic paths from  $X'$  to  $Y'$  according to a balanced almost uniform permutation  $\pi$  on the (signed) weights  $\{w_i\}_{i=1}^m$  of the items in  $S = X' \oplus Y'$ , where  $m = |S|$ .

Stage 3: Send the entire unit flow along a single path from  $Y'$  to  $Y$  by adding the items in  $Y_0$  in index order.

Figure 1 gives a schematic illustration of this flow. Notice how the upper bound in the balance property keeps the path below the hyperplane, while the lower bound keeps the path from passing too close to the origin (thus avoiding a potential bottleneck).

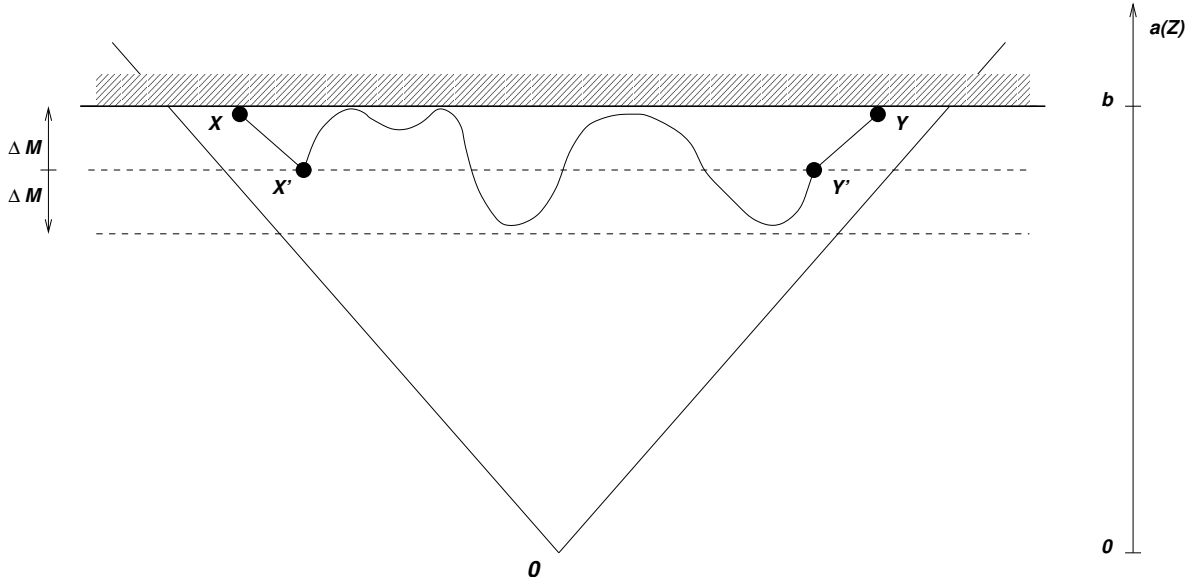


Figure 1: Schematic illustration of flow from  $X$  to  $Y$ . The function  $a(\cdot)$  increases in the vertical direction; the bounding hyperplane is  $a(Z) = b$ . In the case shown,  $a(X') = a(Y')$ ; the balance property of the permutation of  $X' \oplus Y'$  ensures that the path remains within  $\pm\Delta M$  of this level.

Let us first observe that the above flow is valid. For this, we just need to check that all the flow-carrying paths remain within the set  $\Omega$ . This is obvious for stages 1 and 3. For stage 2 it follows from the balance property of  $\pi$ : for if  $Z$  is the  $k$ th point along a flow-carrying path from  $X'$  to  $Y'$ , then

$$\begin{aligned}
 a(Z) &= a(X') + \sum_{i=1}^k w_{\pi(i)} \\
 &\leq a(X') + \max\{a(Y') - a(X'), 0\} + \Delta M \\
 &= \max\{a(Y'), a(X')\} + \Delta M \\
 &\leq b,
 \end{aligned} \tag{16}$$

where in the last line we have used the fact that  $a(X'), a(Y') \leq b - \Delta M$ . Hence  $Z \in \Omega$ .

Next we must bound the quantities  $\mathcal{C}(f)$  and  $\mathcal{L}(f)$  for this flow  $f$ , as defined in section 2.  $\mathcal{L}(f)$ , the length of a longest flow-carrying path, is plainly at most  $m \leq n$ . To estimate the

congestion  $\mathcal{C}(f)$ , we must bound the flow along any edge of  $G_\Omega$ . For convenience in this sketch, we will in fact bound the flow  $f(Z)$  through any *vertex*  $Z$ ; clearly this is also an upper bound on the flow along any edge.

So let  $Z$  be an arbitrary vertex of  $G_\Omega$ . Define  $\mathcal{P}(Z)$  to be the set of pairs  $(X, Y)$  such that some  $X \rightarrow Y$  flow passes through  $Z$ . Note that  $\mathcal{P}(Z) = \bigcup_{i=1}^3 \mathcal{P}_i(Z)$ , where  $\mathcal{P}_i(Z)$  are the pairs whose paths pass through  $Z$  in stage  $i$ . We shall bound the contribution to  $f(Z)$  from each  $\mathcal{P}_i(Z)$  separately. For  $i = 1, 3$  this is simple: since stage-1 paths have length at most  $h$ , for any given  $Y \in G_\Omega$  the number of vertices  $X$  such that  $(X, Y) \in \mathcal{P}_1(Z)$  is (crudely) at most  $(h+1)n^h$ , so the contribution to  $f(Z)$  from such paths is no more than  $(h+1)n^h|\Omega|$ . The same bound holds symmetrically for  $\mathcal{P}_3(Z)$ . The main portion of the paths,  $\mathcal{P}_2(Z)$ , presents more of a challenge.

We shall actually work with  $\mathcal{P}'_2(Z)$ , the set of pairs  $(X', Y')$  such that  $Z$  lies on the stage-2 path with endpoints  $X', Y'$ . By the observation in the previous paragraph, the flow contribution from  $\mathcal{P}_2(Z)$  will be (again crudely) at most  $(h+1)^2 n^{2h}$  times that from  $\mathcal{P}'_2(Z)$ . Recall that we are really interested in the ratio  $\frac{f(Z)}{|\Omega|}$ , rather than in  $f(Z)$  itself. Accordingly, following earlier analyses of this general type (see, e.g., [10, 11]), we measure the set  $\mathcal{P}'_2(Z)$  by associating with each of its elements  $(X', Y')$  an “encoding”  $\tilde{Z}$ , which belongs to  $\Omega$ . This is defined by

$$\tilde{Z} = X' \oplus Y' \oplus Z.$$

To see that  $\tilde{Z} \in \Omega$ , we need to check that  $a(\tilde{Z}) \leq b$ . But this follows because

$$\begin{aligned} a(\tilde{Z}) &= a(X') + a(Y') - a(Z) \\ &\leq a(X') + a(Y') - (\min\{a(X'), a(Y')\} - \Delta M) \\ &= \max\{a(X'), a(Y')\} + \Delta M \\ &\leq b, \end{aligned}$$

where in the second line we have used the balance property of  $\pi$  as in (16) to bound  $a(Z)$ , this time from below.

How many pairs  $(X', Y')$  could be mapped to a given  $\tilde{Z}$ ? First note that  $\tilde{Z}$  uniquely determines both  $S = X' \oplus Y'$  and  $I = X' \cap Y'$  via the relations

$$S = \tilde{Z} \oplus Z; \quad I = \tilde{Z} \cap Z.$$

Thus in particular such pairs share the same symmetric difference,  $S$ , of cardinality  $m$ , say. To determine  $X'$  and  $Y'$  uniquely, it suffices to specify the subset  $U \subseteq S$  of elements that have already been processed (i.e., added or deleted) by the stage-2 path by the time it reaches  $Z$ . For then we know, from the fact that all stage-2 paths are geodesic, that  $X'$  agrees with  $Z$  on  $S - U$  and with  $\tilde{Z}$  on  $U$ , and vice versa for  $Y'$ . More formally,

$$X' = Z \oplus U; \quad Y' = \tilde{Z} \oplus U.$$

The upshot of the foregoing discussion is that we can define a mapping from  $\mathcal{P}'_2(Z)$  to pairs of the form  $(\tilde{Z}, U)$ , where  $\tilde{Z} \in \Omega$  and  $U$  is a subset of  $Z \oplus \tilde{Z}$ . Moreover, and crucially, this mapping is *injective*. It therefore effectively enumerates the set  $\mathcal{P}'_2(Z)$ .

Finally, we need to take account of the actual quantity of flow traveling along the paths. Consider a pair  $(X', Y') \in \mathcal{P}'_2(Z)$ , corresponding to the pair  $(\tilde{Z}, U)$ . Recall that the flow distribution between  $X'$  and  $Y'$  is determined by a balanced almost uniform permutation  $\pi$  of the weights in  $S = X' \oplus Y'$ . The proportion of this flow that passes through  $Z$  is precisely

$$\Pr[\pi\{1, \dots, |U|\} = U] \leq \binom{m}{|U|}^{-1} p(m),$$

by the almost uniform property of  $\pi$ .

Putting all this together, we can bound the total contribution to  $f(Z)$  from  $\mathcal{P}'_2(Z)$  as follows:

$$\begin{aligned} \sum_{\tilde{Z} \in \Omega} \sum_{U \subseteq Z \oplus \tilde{Z}} \Pr[\pi\{1, \dots, |U|\} = U] &\leq \sum_{\tilde{Z} \in \Omega} \sum_k \sum_{U \subseteq Z \oplus \tilde{Z}, |U|=k} \binom{m}{k}^{-1} p(m) \\ &\leq p(n) \sum_{\tilde{Z} \in \Omega} \sum_k \binom{m}{k} \binom{m}{k}^{-1} \\ &\leq np(n)|\Omega|, \end{aligned}$$

where in the summations  $m = |Z \oplus \tilde{Z}|$ . The total contribution from all stage-2 paths is thus at most  $(h+1)^2 n^{2h+1} p(n) |\Omega|$ .

Combining this with our earlier bounds for stages 1 and 3, we obtain that  $f(Z) \leq (2(h+1)n^h + (h+1)^2 n^{2h+1} p(n)) |\Omega|$ , and hence  $\mathcal{C}(f) \leq p'(n)$  (for a different polynomial  $p'$ ). Since both  $\mathcal{L}(f)$  and  $\mathcal{C}(f)$  are bounded polynomially in  $n$ , we obtain immediately from Theorem 2.1 that the mixing time,  $\tau_{\text{mix}}$ , is polynomial in  $n$  as required.

In the above sketch, we assumed that we could remove a fixed number of items from  $X$  and  $Y$  and thereby ensure that they are well separated from the bounding hyperplane. The main extra work in the full description of the flow in the next subsection is concerned with discharging this assumption. In addition, we will be rather more precise in our accounting so as to achieve the tightest possible bound on  $\mathcal{C}(f)$ .

## 4.2 Complete description

In this subsection, we present a complete construction, for an arbitrary instance of the 0-1 knapsack problem, of a flow  $f$  with congestion  $\mathcal{C}(f) = O(n^{3/2+\epsilon})$  for any  $\epsilon > 0$  and length  $\mathcal{L}(f) = O(n)$ . By Theorem 2.1, this will imply that the mixing time of the random walk on  $G_\Omega$  is  $O(n^{9/2+\epsilon})$ . From now on we assume that  $\epsilon > 0$  is arbitrary but fixed.

Let  $X, Y$  be arbitrary vertices of  $G_\Omega$ . Recall the scheme for constructing a flow from  $X$  to  $Y$  in the sketch of section 4.1: we essentially followed a balanced almost uniform permutation of  $X \oplus Y$ , except that we removed a constant number of items from consideration (processing them at the beginning and end of the path) to ensure that both path endpoints were far enough from the hyperplane so that the path remained within  $G_\Omega$ . In reality this property is not guaranteed by the removal of a fixed number of items because of the possibly large variations in weights between  $X$  and  $Y$ . (For example, if all the items in  $X$  are much larger than those in  $Y$ , then removing any fixed number of items from  $X$  and  $Y$  could still leave  $Y$  too close to the hyperplane relative to the size of the largest remaining item in  $X \oplus Y$ .) However, it turns out that after removing a fixed number of heavy items from  $X \oplus Y$ , we can perform a “pre-processing” operation by randomly switching items between  $X$  and  $Y$  to roughly balance their weights. Moreover, we will need to add and delete the removed items repeatedly along the path to maintain fine balance. The resulting flow-carrying paths will not in general be geodesics, as before, though they will have length only  $O(n)$ .

In preparation for describing the flow, we first describe the pre-processing operation. We assume that  $a(X) + a(Y) \leq 2b - 6\Delta M$ , where  $M = \max_{i \in X \oplus Y} a_i$  and  $\Delta = \Delta(\epsilon)$  is the constant appearing in Theorem 3.4. (We will reduce to this case by first deleting a fixed number of items from  $X \oplus Y$ .) Call the pair  $(X, Y)$  *full* if either  $a(X) > b - \Delta M$  or  $a(Y) > b - \Delta M$ . Our goal is to shift items randomly between  $X$  and  $Y$  and thereby reach a pair  $(X', Y')$  that is not full.

Consider the following random walk on  $\{(X', Y') : X' \cup Y' = X \cup Y, X' \cap Y' = X \cap Y, a(X'), a(Y') \leq b\}$ . If the current state is  $(X', Y')$ , choose an index  $i \in X' \oplus Y'$  u.a.r. With probability  $\frac{1}{2}$ , do nothing; else move  $i$  from  $X'$  to  $Y'$  or  $Y'$  to  $X'$  if possible. We call this the

“pre-processing random walk” (PRW). We claim in the following lemma that, if we run the PRW for a number of steps chosen randomly between 1 and  $O(n)$ , we will with reasonable probability arrive at a pair  $(X', Y')$  that is not full. The proof uses a martingale argument and is deferred to section 6.

**Lemma 4.1** *Let  $(X, Y)$  be a full pair of vertices in  $G_\Omega$  with  $a(X) + a(Y) \leq 2b - 6\Delta M$ , where  $M = \max_{i \in X \oplus Y} a_i$ . Pick  $T$  u.a.r. from  $\{1, 2, \dots, C_1 m\}$ , where  $m = |X \oplus Y|$  and  $C_1$  is a suitable constant (which depends only on  $\Delta$ ), and let  $(X', Y')$  be the result of running the PRW for  $T$  steps starting from  $(X, Y)$ . Then  $\Pr[(X', Y') \text{ is not full}] \geq 1/C_2$  for a positive constant  $C_2$  (again depending only on  $\Delta$ ).*

We are now ready to construct and analyze the flow.

**Lemma 4.2** *For arbitrary weights and any  $\epsilon > 0$ , it is possible to construct a multicommodity flow  $f$  in  $G_\Omega$  with  $\mathcal{C}(f) = O(n^{3/2+\epsilon})$  and  $\mathcal{L}(f) = O(n)$ .*

**Proof:** Let  $X, Y$  be arbitrary vertices of  $G_\Omega$ . Viewing  $X$  and  $Y$  as subsets of  $\{1, \dots, n\}$ , let  $H$  be the  $h = \lceil 6\Delta \rceil$  elements of  $X \oplus Y$  having largest weight (or let  $H = X \oplus Y$  if  $|X \oplus Y| \leq h$ ), with ties broken according to index order. Define  $X' = X - H$ ,  $Y' = Y - H$  and  $S = X' \oplus Y'$ . Let  $m = |S|$ , and let  $\{w_i\}_{i=1}^m$  be an arbitrary enumeration of the weights of the items in  $S$ , with the weights of items in  $X', Y'$  appearing with negative and positive signs respectively. Let  $M = \max_i |w_i|$ .

We will say that a set of indices  $Z$  is *good* if  $Z - H \in \Omega$  and  $(Z \oplus X \oplus Y) - H \in \Omega$ . For a set of indices  $Z$  and an index  $i$ , define

$$Zi = \begin{cases} Z \oplus \{i\} & \text{if } Z \oplus \{i\} \text{ is good;} \\ Z & \text{otherwise.} \end{cases}$$

Define  $Zi_1i_2 = ((Zi_1)i_2)$  and so on. Note that if  $Y = Xi_1 \cdots i_l$  then the sequence  $i_1, \dots, i_l$  defines a path from  $X$  to  $Y$  in the unit hypercube of length at most  $l$ . This path need not in general lie within  $G_\Omega$ ; however, it is “close to”  $G_\Omega$  in the sense that for every point  $Z$  of the path,  $Z - H \in \Omega$ .

If  $(X', Y')$  is not full then set  $T = 0$ , otherwise choose  $T$  u.a.r. from  $\{1, \dots, C_1 m\}$ , where  $C_1$  is the constant in Lemma 4.1. Next, let  $i_1, \dots, i_T$  be i.i.d. uniform over  $S$ . Define  $X'' = X'i_1 \cdots i_T$  and  $Y'' = X'' \oplus X' \oplus Y' = Y'i_1 \cdots i_T$ . Thus  $(X'', Y'')$  is the result of running the PRW for  $T$  steps starting from  $(X', Y')$ . Note that  $a(X') + a(Y') \leq 2b - a(H) \leq 2b - 6\Delta M$ . So, by Lemma 4.1, we can condition on the event that the pair  $(X'', Y'')$  is not full and thus increase the probability of any path by a factor of at most  $C_2$ .

Now let  $\pi$  be a  $\Delta$ -balanced, strongly  $Cm^{1/2+\epsilon}$ -uniform permutation on the weights  $\{w_i\}$ , whose existence is guaranteed by Theorem 3.4. We claim that the sequence

$$i_1, \dots, i_T, \pi(1), \dots, \pi(m), i_T, \dots, i_1 \tag{17}$$

defines a path from  $X$  to  $Y$  in the hypercube. This is true because the condition that  $(X'', Y'')$  be not full, together with the fact that  $\pi$  is balanced, guarantees that all of the transitions indicated by  $\pi$  will actually take place.

Set

$$j_k = \begin{cases} i_k & \text{if } 1 \leq k \leq T; \\ \pi(k - T) & \text{if } T < k \leq T + m; \\ i_{2T+m-k-1} & \text{if } T + m < k \leq 2T + m, \end{cases}$$

and let  $l = 2T + m$ . Then  $j_1, \dots, j_l$  is the sequence in (17). Our flow from  $X$  to  $Y$  will essentially follow the sequence  $j_k$ , except that along the way elements of  $H$  will be used to keep the knapsack as full as possible, but will be removed as necessary to make room for new items  $j_k$  to be added. Thus each intermediate state  $Z$  will be of the form  $\overline{H} \oplus Xj_1 \cdots j_k$ , for some  $\overline{H} \subseteq H$  and  $k \leq l$ .

Suppose that, after processing the first  $k \leq l$  elements of the sequence in (17), we have  $Z = \overline{H} \oplus Xj_1 \cdots j_k$  for some  $\overline{H} \subseteq H$ . The transition rule will be as follows.

1. If  $k < l$  and  $j_{k+1} \notin Z$  then move to  $Z \oplus j_{k+1}$  if possible (i.e., if the result is an element of  $\Omega$ ); otherwise delete an element from  $H$ .
2. If  $k < l$  and  $j_{k+1} \in Z$  then add an element from  $H$  if possible; otherwise move to  $Z \oplus j_{k+1}$ .
3. If  $k = l$  then add an element from  $H \cap Y$  if possible; otherwise delete an element from  $H \cap X$ .

The fact that all of the sets  $Xj_1 \cdots j_k$  are good ensures that sufficient elements of  $H$  can always be removed so as to make room to add the next element  $j_{k+1}$  when necessary; hence the above rule defines a feasible random path from  $X$  to  $Y$ . Similarly, goodness also implies that  $a(Z \oplus X \oplus Y - H) \leq b$  for every intermediate state  $Z$ ; since our rule keeps the weight as large as possible this implies that, at any intermediate edge  $(Z, W)$  along the path, there exists (at most) one element  $u \in H$  such that

$$a(Z \oplus X \oplus Y - \{u, z\}) \leq b, \quad (18)$$

where  $z$  is the index such that  $\{z\} = Z \oplus W$ . Then  $(\tilde{Z} - \{u, z\}) \in \Omega$ , where exactly as in the analysis in section 4, we define the ‘‘encoding’’  $\tilde{Z}$  by

$$\tilde{Z} = X \oplus Y \oplus Z.$$

Thus, for any given edge  $(Z, W)$ , the number of encodings  $\tilde{Z}$  is at most  $n|\Omega|$ .

Note that the path from  $X$  to  $Y$  can be naturally divided into three *stages*, corresponding to the three parts of the sequence  $j_k$ . We will write the flow through any given edge  $(Z, W) \in G_\Omega$  as  $f(Z, W) = f_1(Z, W) + f_2(Z, W) + f_3(Z, W)$ , where  $f_i(Z, W)$  is the contribution of stage  $i$  paths. We will bound  $f$  by bounding each of the three contributions  $f_i$  separately.

Consider stage 1 first, and focus on a particular edge  $(Z, W)$ . For any pair  $(X, Y)$  that sends flow through  $(Z, W)$  in stage 1, we can write  $Z = \overline{H} \oplus Xj_1 \cdots j_k$ , where  $j_1, \dots, j_k$  are the first  $k$  elements processed along the path. Thus the pair  $(X, Y)$  is completely specified by  $k, j_1, \dots, j_k, \tilde{Z}$  and  $\overline{H}$ , via the easily verified relations

$$X = \overline{H} \oplus Zj_k \cdots j_1; \quad Y = \overline{H} \oplus \tilde{Z}j_k \cdots j_1.$$

The amount of flow corresponding to a given sequence  $j_1, \dots, j_k$  is bounded above by the probability that  $j_1, \dots, j_k, z$  were the first  $k + 1$  indices chosen in the pre-processing random walk, which is at most  $C_2 m^{-(k+1)}$ . (The factor  $C_2$  here arises from our earlier conditioning on the event that  $(X'', Y'')$  is not full.) Thus we can bound the stage-1 flow  $f_1(Z, W)$  as in section 4. We have

$$\begin{aligned} f_1(Z, W) &\leq \sum_{\tilde{Z}} \sum_k \sum_{j_1, \dots, j_k} \sum_{\overline{H}} C_2 m^{-(k+1)} \\ &\leq \sum_{\tilde{Z}} \sum_k 2^h C_2 m^{-1} \\ &\leq \sum_{\tilde{Z}} C_1 m 2^h C_2 m^{-1} \\ &\leq C_1 C_2 2^h n |\Omega|, \end{aligned}$$

where the factors  $C_1 m$  and  $2^h$  arise from summing over  $k$  and  $\overline{H}$  respectively.

The flow  $f_3(Z, W)$  from stage-3 paths can be handled symmetrically, so consider now the stage-2 paths. For a given edge  $(Z, W)$ , every pair  $(X, Y)$  that sends flow through  $(Z, W)$  in stage 2 can be completely specified by  $\tilde{Z}, T, k, j_1, \dots, j_T, U$  and  $\overline{H}$ , where  $k$  is the number of elements of the sequence in (17) processed along the path from  $X$  to  $Z$  and  $U = \{\pi(1), \dots, \pi(k - T)\}$ , via

$$X = \overline{H} \oplus (Z \oplus U)j_T \cdots j_1; \quad Y = \overline{H} \oplus (\tilde{Z} \oplus U)j_T \cdots j_1.$$

Let  $k' = k - T$ . The amount of flow corresponding to a given  $j_1, \dots, j_T$  and  $U$  is bounded above by

$$(C_2 m^{-T})(C_1 m)^{-1} \left[ C m^{1/2+\epsilon} \binom{m}{k', m - k' - 1, 1}^{-1} \right],$$

where the first factor comes from the pre-processing random walk, the second factor is the probability of choosing a particular  $T$ , and the third factor is an upper bound on the probability  $\Pr[\pi\{1, \dots, k'\} = U \text{ and } \pi(k' + 1) = z]$ , which comes from the strong almost uniformity of  $\pi$ . Thus we can again bound the flow  $f_2(Z, W)$  as in section 4. We have

$$\begin{aligned} f_2(Z, W) &\leq \sum_{\tilde{Z}} \sum_T \sum_k \sum_{j_1, \dots, j_T} \sum_U \sum_{\overline{H}} (C_2 m^{-T})(C_1 m)^{-1} \left[ C m^{1/2+\epsilon} \binom{m}{k', m - k' - 1, 1}^{-1} \right] \\ &\leq \sum_{\tilde{Z}} (C_1 m) m m^T \binom{m-1}{k'} 2^h (C_2 m^{-T})(C_1 m)^{-1} \left[ C m^{1/2+\epsilon} \binom{m}{k', m - k' - 1, 1}^{-1} \right] \\ &= \sum_{\tilde{Z}} 2^h C_2 \left[ m \binom{m-1}{k'} \binom{m}{k', m - k' - 1, 1}^{-1} \right] C m^{1/2+\epsilon} \\ &= \sum_{\tilde{Z}} 2^h C_2 C m^{1/2+\epsilon} \\ &\leq 2^h C C_2 n^{3/2+\epsilon} |\Omega|, \end{aligned}$$

where the factors in the second line are written in the same order as the sums they arise from.

Adding the contributions  $f_1, f_2$  and  $f_3$ , we see that the above flow satisfies  $\mathcal{C}(f) = O(n^{3/2+\epsilon})$ , while plainly  $\mathcal{L}(f) = O(n)$ . Since  $\epsilon > 0$  was arbitrary, this completes the proof.  $\square$

Given such a flow, we need only invoke Theorem 2.1 to derive our main result:

**Theorem 4.3** *Let  $\Omega$  be the set of solutions to an arbitrary instance of the 0-1 knapsack problem. The mixing time of the random walk on  $G_\Omega$  is  $\tau_{\text{mix}} = O(n^{9/2+\epsilon})$  for any  $\epsilon > 0$ .*

As mentioned in the Introduction, this immediately yields an fpras for computing  $|\Omega|$ , via a standard reduction to random sampling (whose details are spelled out in [12]).

**Remark:** The mixing time bound of  $O(n^{9/2+\epsilon})$  in Theorem 4.3 is reasonably tight for this type of analysis. If we apply Theorem 2.1 to analyze random walk on the entire cube  $\{0, 1\}^n$ , we get a bound of  $O(n^3)$  even with an optimal flow. Thus the truncation introduces an extra factor of only  $O(n^{3/2+\epsilon})$  into the bound. It is instructive to see where this extra factor comes from:  $O(n^{1/2+\epsilon})$  is due to the balanced almost uniform permutation construction (Theorem 3.4, which is tight), while  $O(n)$  comes from the fact that the “encoding”  $\tilde{Z}$  may lie just outside  $\Omega$ .  $\square$

## 5 Multiple hyperplanes

### 5.1 Introduction

In this section, we will extend our earlier results to handle multiple hyperplanes. For a non-negative real  $d \times n$  matrix  $A = (a_{ij})$  and a positive real vector  $\mathbf{b} = (b^1, \dots, b^d)$ , let  $\Omega$  denote the set of 0-1 vectors  $\mathbf{x} = (x_i)_{i=1}^n$  for which  $\mathbf{A}\mathbf{x} \leq \mathbf{b}$ . The vertices in  $\Omega$  constitute the set of feasible solutions to the multidimensional knapsack problem with the  $d$  simultaneous constraints

$$\mathbf{a}^j \cdot \mathbf{x} \equiv \sum_{i=1}^n a_i^j x_i \leq b^j \quad \text{for } 1 \leq j \leq d, \quad (19)$$

where  $a_i^j \equiv a_{ji}$ . (In equation (19) the superscript  $j$  indexes the  $j$ th linear constraint; we will follow this convention throughout.)

Geometrically,  $\Omega$  is obtained by truncating the unit cube by  $d$  hyperplanes, each of which corresponds to a knapsack constraint. The essential geometric property of these “knapsack” hyperplanes is that their normal vectors all lie in the same quadrant. The results of this section will easily extend to any collection of hyperplanes with this property<sup>7</sup>.

Following our earlier notation, we identify a 0-1 vector  $\mathbf{x} = (x_i)_{i=1}^n$  with the set of indices  $X = \{i : x_i = 1\}$ , and write  $a(X) = (a^1(X), \dots, a^d(X))$  for the (now  $d$ -dimensional) weight of  $X$ . As before we denote by  $G_\Omega$  the subgraph of the hypercube  $\{0, 1\}^n$  induced by the vertices in  $\Omega$ , and we again study symmetric random walk on  $G_\Omega$ ; i.e., transitions from a given state  $X \subseteq \{1, \dots, n\}$  are made as follows:

1. pick an item  $i \in \{1, \dots, n\}$  u.a.r.;
2. if  $i \in X$ , move to  $X - \{i\}$ ; if  $i \notin X$  and  $a^j(X \cup \{i\}) \leq b^j$  for all  $j$ , move to  $X \cup \{i\}$ ; otherwise, do nothing.

Again, to avoid issues involving periodicity, we add to every state a holding probability of  $\frac{1}{2}$ .

In this section we will prove that this random walk on  $G_\Omega$  has mixing time that is polynomially bounded in  $n$ , for any fixed dimension  $d$ . Just as in the one-dimensional case, this immediately gives a polynomial time algorithm for sampling (almost) uniformly at random from  $\Omega$ , and a fpras for computing  $|\Omega|$ .

We note that the degree of our polynomial upper bound for the mixing time will depend on the dimension  $d$ , but this is unavoidable as the following simple example shows. Consider a  $d$ -dimensional knapsack problem in which there are  $\frac{n}{2d}$  items having each of the  $d$  weight vectors  $(n, 0, \dots, 0)$ ,  $(0, n, \dots, 0)$ ,  $\dots$ ,  $(0, \dots, 0, n)$ , and the remaining  $\frac{n}{2}$  items have weight vector  $(1, 1, \dots, 1)$ ; the knapsack capacity is  $\mathbf{b} = (n, \dots, n)$ . Let  $S$  be the set of feasible solutions in  $\Omega$  which do not contain any of the  $(1, \dots, 1)$  items. Then  $|S| = (\frac{n}{2d} + 1)^d$ , but  $S$  is connected to  $\Omega - S$  only through the origin. It follows easily that the mixing time is  $n^{\Omega(d)}$ .

In fact, for arbitrary  $d$  there can be no uniform polynomial upper bound for the running time of *any* sampling algorithm unless  $\text{RP} = \text{NP}$ . This follows immediately by reduction from the problem of sampling independent sets in a graph. By theorem 1.17 of [19], there is no algorithm for (almost) uniformly sampling independent sets in a graph unless  $\text{RP} = \text{NP}$ . Now if  $G = (V, E)$  is an arbitrary (undirected) graph, there is a 1-1 correspondence between the independent sets in

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<sup>7</sup>However, we cannot allow the hyperplanes to be arbitrary. If arbitrary truncations were allowed, then it would be possible to use just two hyperplanes to cause exponential bottlenecks or even disconnect the graph  $G_\Omega$ .

$G$  and the feasible solutions to the knapsack problem with  $|V|$  variables and the  $|E|$  constraints  $x_u + x_v \leq 1$  for all  $\{u, v\} \in E$ .

To prove rapid mixing of the random walk on  $G_\Omega$  for any fixed  $d$ , we use the multicommodity flow technique as before. Recall that Theorem 2.1, which bounds the mixing time in terms of the cost of a flow  $f$ , holds for symmetric random walk on any connected subset of the hypercube, so it again suffices to come up with a flow of small cost. As before, the idea is to spread each  $X \rightarrow Y$  flow evenly using a balanced almost uniform permutation. However, since the weight function  $a(\cdot)$  is now vector-valued, we first need to extend the definition of balance to higher dimensions.

**Definition 5.1** Fix an integer  $d > 0$ , and let  $\{w_i\}_{i=1}^m$  be a set of weights in  $\mathbf{R}^d$  satisfying  $\sum_{i=1}^m w_i = 0$ . For a positive real number  $\Delta$ , a permutation  $\pi \in \mathcal{S}_m$  is  $\Delta$ -balanced if

$$\max_k \left| \sum_{i=1}^k w_{\pi(i)}^j \right| \leq \Delta M^j \quad \text{for } 1 \leq j \leq d, \quad (20)$$

where  $w_i = (w_i^1, \dots, w_i^d)$  and  $M^j = \max_{1 \leq i \leq m} |w_i^j|$ .

Thus  $\pi$  is balanced with respect to vector weights  $\{w_i\}$  if and only if it satisfies the  $d$  one-dimensional balance conditions given by (20). Note that this generalizes our earlier Definition 3.1 for the one-dimensional case (except that, for simplicity, we have assumed that  $\sum_i w_i = 0$ ).

Constructing balanced almost uniform permutations is significantly more difficult in higher dimensions since one has to control fluctuations in all dimensions simultaneously. In fact, for  $d \geq 2$ , it is non-trivial to prove for an arbitrary set of vector weights that even a *single* balanced permutation exists. (For  $d = 1$  of course this is trivial.) The existence of such a permutation follows at once from a lemma due to Grinberg and Sevast'yanov [7], which was proved in an entirely different context:

**Lemma 5.2** [7] Let  $x_1, \dots, x_n$  be vectors in  $\mathbf{R}^d$  such that  $\sum_i x_i = 0$ . Then there exists a permutation  $\nu \in \mathcal{S}_n$  such that

$$\sum_{i=1}^k x_{\nu(i)} \in d \times \text{conv}\{x_1, \dots, x_n\} \quad \text{for } 1 \leq k \leq n,$$

where  $d \times \text{conv}$  denotes the dilation of the convex hull by  $d$ .

Of course, we need something much stronger than this, namely *almost uniform* permutations with a similar balance property. Perhaps surprisingly, we will show that balanced almost uniform permutations exist in arbitrary dimension  $d$ . To illustrate the main ideas involved in extending from one to higher dimensions, we now give a sketch of the proof in the special case where  $d = 2$  and the weights satisfy  $1 \leq |w_i^j| \leq 2$  for all  $i$  and  $j$ .

In this setting, let  $I_1 = \{i : w_i^2 \geq 0\}$ ,  $I_2 = \{i : w_i^2 < 0\}$ , and define  $v = \sum_{i \in I_1} w_i$ . For every  $i \leq m$ , let  $y_i$  be the projection of  $w_i$  onto  $v^\perp$ . Let  $\pi_1$  be an almost uniform permutation on  $I_1$  which is balanced (in the one-dimensional sense) with respect to  $\{y_i\}_{i \in I_1}$ , with a similar definition for  $\pi_2$ . Finally, interleave  $\pi_1$  and  $\pi_2$  to give a permutation on  $\{1, \dots, m\}$  which is balanced with respect to  $\{w_i^2\}_{i=1}^m$  (the projections of the  $w_i$  onto the second coordinate axis). Since  $\pi_1$  and  $\pi_2$  are both almost uniform, so is  $\pi$ , by an argument similar to that in the proof of Lemma 3.3.

Furthermore, since  $\pi_1$  and  $\pi_2$  are each balanced with respect to projections onto  $v^\perp$ , so is  $\pi$ . (Note that the projections  $y_i$  satisfy  $\sum_{i \in I_1} y_i = \sum_{i \in I_2} y_i = 0$ .) Thus, for every  $k$ , the projections of  $\sum_{i=1}^k w_{\pi(i)}$  onto the second coordinate axis and onto  $v^\perp$  are both bounded, and since the  $w_i^j$  are

all in  $[1, 2]$ , the angle between the coordinate axis and  $v^\perp$  is bounded away from zero. Thus, the partial sums  $\sum_{i=1}^k w_{\pi(i)}$  stay inside a parallelogram of bounded diameter. Hence  $\pi$  is balanced with respect to the weights  $\{w_i\}_{i=1}^m$ .

This concludes the sketch proof for the above special case with  $d = 2$ . Note that it is a straightforward reduction to the one-dimensional result. Unfortunately, in general the reduction from  $d$  to  $d-1$  dimensions is not quite so straightforward; we deal with the extra technical difficulties in the next subsection.

## 5.2 Balanced almost uniform permutations in arbitrary dimensions

The following theorem says that one can always construct balanced almost uniform permutations when the dimension  $d$  is fixed.

**Theorem 5.3** *Let  $d$  be any positive integer. There is a constant  $c_d$  and a polynomial function  $p_d$  such that, for any set of weights  $\{w_i\}_{i=1}^m$  in  $\mathbf{R}^d$  with  $\sum_i w_i = 0$ , there exists a  $c_d$ -balanced,  $p_d(m)$ -uniform permutation.*

**Proof:** The proof will be by induction on  $d$ . The base case  $d = 1$  follows from Corollary 3.5, with  $c_1 = 9$  and  $p_1(m) = Cm^2$ . Now let  $d > 1$  be arbitrary, and suppose that the result holds for dimensions up to  $d-1$ . Let  $\{w_i\}_{i=1}^m$  be a set of weights in  $\mathbf{R}^d$ . Suppose first that the weights satisfy

$$M^j = 2 \quad \text{for all } j; \quad (21)$$

$$1 \leq \max_{1 \leq j \leq d} |w_i^j| \leq 2 \quad \text{for all } i. \quad (22)$$

Thus each weight is at least half as large as the maximum (positive or negative) weight in some coordinate. Then

$$\max_{1 \leq j \leq d} \sum_{i=1}^m |w_i^j| \geq \frac{m}{d}.$$

W.l.o.g., suppose that the sum in the LHS is maximized by  $j = d$ . Then we have

$$\sum_{i=1}^m (w_i^d)^+ = \sum_{i=1}^m (w_i^d)^- \geq \frac{m}{2d},$$

where we are using the notation  $z^+ = \max\{z, 0\}$  and  $z^- = \max\{-z, 0\}$ .

Let  $I_1 = \{i : w_i^d \geq 0\}$ ,  $I_2 = \{i : w_i^d < 0\}$ ,  $m_1 = |I_1|$ , and  $m_2 = |I_2|$ . Define the means  $\mu_1 = \frac{1}{m_1} \sum_{i \in I_1} w_i^d$ , and  $\mu_2 = -\frac{1}{m_2} \sum_{i \in I_2} w_i^d$ . Note that  $\mu_1, \mu_2 \geq \frac{1}{2d}$ . For  $1 \leq j < d$ , let

$$\gamma^j = \frac{\sum_{i \in I_1} w_i^j}{\sum_{i \in I_1} w_i^d} = \frac{\sum_{i \in I_2} w_i^j}{\sum_{i \in I_2} w_i^d}.$$

For all  $i \leq m$  and  $j < d$ , let  $y_i^j = w_i^j - \gamma^j w_i^d$ , and let  $y_i = (y_i^1, \dots, y_i^{d-1})$ . Note that  $|\gamma^j| \leq 1$ ,  $|y_i^j| \leq 4$ , and  $\sum_{i \in I_1} y_i = \sum_{i \in I_2} y_i = 0$ .

Now, for  $s = 1, 2$  let  $\pi_s$  be a  $p_{d-1}(m)$ -uniform permutation on  $I_s$  which is  $c_{d-1}$ -balanced with respect to  $\{y_i\}_{i \in I_s}$ . Call  $\pi_1$   $\alpha$ -good if for every  $k_1$  with  $1 \leq k_1 \leq m_1$  we have

$$\left| \sum_{i=1}^{k_1} w_{\pi_1(i)}^d - k_1 \mu_1 \right| \leq 2\alpha \sqrt{k_1^*} \quad (23)$$

where  $k_1^* = \min\{k_1, m_1 - k_1\}$ . In similar fashion to the proof of Lemma 6.1, Hoeffding's bounds [8] imply that for a particular value of  $k_1$  we have

$$\Pr_{\text{unif}}(\pi_1 \text{ does not satisfy (23)}) \leq 2 \exp(-2\alpha^2),$$

and since the event depends only on the initial segment  $\pi_1\{1, \dots, k_1\}$ , we also have

$$\begin{aligned} \Pr(\pi_1 \text{ does not satisfy (23)}) &\leq p_{d-1}(m) \cdot \Pr_{\text{unif}}(\pi_1 \text{ does not satisfy (23)}) \\ &\leq p_{d-1}(m) \cdot 2 \exp(-2\alpha^2). \end{aligned}$$

Hence

$$\Pr[\pi_1 \text{ is not } \alpha\text{-good}] \leq mp_{d-1}(m) \cdot 2 \exp(-2\alpha^2). \quad (24)$$

Suppose that for some constants  $C$  and  $r$ , the polynomial  $p_{d-1}$  satisfies  $p_{d-1}(k) \leq Ck^r$  for all  $k$ . If we let  $\alpha = \sqrt{(r+1)\ln(m)}$ , then the RHS of (24) is at most  $2Cm^{r+1-2(r+1)} \leq \frac{1}{2}$ , for sufficiently large  $m$ . Thus, we can assume that  $\pi_1$  is  $\alpha$ -good with probability 1 and only increase  $C$  by a constant factor. Similar arguments apply to  $\pi_2$ .

Finally, note that it is always possible to interleave  $\pi_1$  and  $\pi_2$  to give a permutation on  $\{1, \dots, m\}$  which is 1-balanced with respect to  $\{w_i^d\}_{i=1}^m$ . Let  $\pi$  be such a permutation. Then we have  $|\sum_{i=1}^k w_{\pi(i)}^d| \leq 2$ , and

$$\begin{aligned} \left| \sum_{i=1}^k w_{\pi(i)}^j \right| &= \left| \sum_{\substack{i \in I_1: \\ i \leq k}} w_{\pi(i)}^j + \sum_{\substack{i \in I_2: \\ i \leq k}} w_{\pi(i)}^j \right| \\ &= \left| \sum_{\substack{i \in I_1: \\ i \leq k}} y_{\pi(i)}^j + \sum_{\substack{i \in I_2: \\ i \leq k}} y_{\pi(i)}^j + \gamma^j \sum_{i=1}^k w_{\pi(i)}^d \right| \\ &\leq \left| \sum_{\substack{i \in I_1: \\ i \leq k}} y_{\pi(i)}^j \right| + \left| \sum_{\substack{i \in I_2: \\ i \leq k}} y_{\pi(i)}^j \right| + |\gamma^j| \left| \sum_{i=1}^k w_{\pi(i)}^d \right| \\ &\leq 4c_{d-1} + 4c_{d-1} + 2|\gamma^j| \\ &\leq 8c_{d-1} + 2, \end{aligned}$$

for all  $j < d$  and  $k$ . Hence  $\pi$  is  $c'_d$ -balanced for  $c'_d = 4c_{d-1} + 1$  by assumption (21).

To verify almost uniformity, we follow the proof of Lemma 3.3. Let  $U \subseteq \{1, \dots, m\}$  be arbitrary with  $|U| = k$ , and let  $U_1 = U \cap I_1$ ,  $U_2 = U \cap I_2$ ,  $k_1 = |U_1|$ , and  $k_2 = |U_2|$ . Then we have

$$\begin{aligned} \Pr[\pi\{1, \dots, k\} = U] &\leq \Pr[\pi_1\{1, \dots, k_1\} = U_1 \text{ and } \pi_2\{1, \dots, k_2\} = U_2] \\ &\leq (Cm^r)^2 \Pr_{\text{unif}}[\pi_1\{1, \dots, k_1\} = U_1 \text{ and } \pi_2\{1, \dots, k_2\} = U_2] \\ &= \frac{C^2 m^{2r}}{\binom{m_1}{k_1} \binom{m_2}{k_2}} \end{aligned}$$

Now we can bound the quantity  $\frac{\binom{m_1}{k_1} \binom{m_2}{k_2}}{C^2 m^{2r}}$  by mimicking (with minor modifications) the calculations from equation (7) to equation (10) in the proof of Lemma 3.3. In our current setting, we have  $\mu_1, \mu_2 \geq \frac{1}{2d}$ , and the  $|w_i^d|$  are in  $[0, 2]$ . Because we have changed the definition of  $\alpha$ -good and the value of  $\alpha$ , we also have to make the substitutions  $(B-1)^2 \rightarrow 2^2$  and  $\ln m \rightarrow (r+1)\ln m$ , respectively. Thus the bound on the exponent given in equation (9) becomes

$$\frac{10 \cdot 2^2}{\left(\frac{1}{2d}\right)^2} (r+1) \ln m = 160d^2 (r+1) \ln m. \quad (25)$$

Hence  $\pi$  is  $p_d(m)$ -uniform for  $p_d(m) = C'' m^{160d^2(r+1)+2r+1/2}$ .

We have shown how to make balanced almost uniform permutations if the weights satisfy (21) and (22). To generalize to arbitrary weights  $\{w_i\}_{i=1}^m$ , we use the interval trick introduced in section 3.2. Let  $\beta$  be a uniform random permutation in  $\mathcal{S}_m$ , and let  $T_1 = \min\{t : |\sum_{i=1}^t w_i^j| > M^j \text{ for some } j\}$ . Define  $T_2, T_3, \dots$  similarly. Now use the  $T_i$  to divide  $\beta$  into intervals  $I_1, \dots, I_q$ . Let  $\{\alpha_i\}_{i=1}^{q-1}$  be the aggregated ( $d$ -dimensional) weights of the first  $q-1$  intervals. Note that if we divide each  $\alpha_i^j$  by  $\frac{1}{2} \max_i |\alpha_i^j|$ , then the resulting weights satisfy (21) and (22). Hence these weights admit a  $c'_d$ -balanced,  $p_d(q)$ -uniform permutation (though see the remark immediately following this proof). Rearranging the intervals  $\{I_i\}_{i=1}^{q-1}$  according to such a permutation gives a permutation on  $\{1, \dots, m\}$  which is  $p_d(m)$ -uniform and  $c_d$ -balanced for  $c_d = 2c'_d + 1$ .  $\square$

**Remark:** We should point out that the weights  $\{\alpha_i\}_{i=1}^{q-1}$  of the first  $q-1$  intervals will not in general sum to zero. However, we can easily get around this by introducing a dummy weight  $\alpha_q$  which is equal to the weight of interval  $I_q$ . The presence of this single small weight does not affect equation (25) for sufficiently large  $m$ . Hence there is a  $c'_d$ -balanced  $p_d$ -uniform permutation on this padded sequence  $\{\alpha_i\}_{i=1}^q$ . This induces a permutation on  $\{\alpha_i\}_{i=1}^{q-1}$  which is  $(c'_d + 1)$ -balanced and  $cp_d(q)$ -uniform for some constant  $c$ . Thus, if we incorporate an extra  $+1$  into the constant  $c'_d$  and an extra factor of  $c$  into  $p_d$  then the argument in the above proof is still valid.  $\square$

Before we specify our flow, we need one more definition.

**Definition 5.4** Let  $\{w_i\}_{i=1}^m$  be a sequence in  $\mathbf{R}^d$ , with  $w_i = (w_i^1, \dots, w_i^d)$ , let  $\mu = (\mu^1, \dots, \mu^d) = \frac{1}{m} \sum_{i=1}^m w_i$ , and let  $l$  be a positive integer. A permutation  $\pi$  is strongly  $l$ -balanced if, for all  $k \leq m$  and  $j \leq d$ , there exists a set  $S \subseteq \{1, \dots, m\}$  with  $|S \oplus \pi\{1, \dots, k\}| \leq l$ , such that  $(\sum_{i=1}^k w_{\pi(i)}^j - k\mu^j)$  and  $(\sum_{i \in S} w_{\pi(i)}^j - k\mu^j)$  have opposite signs (or either is 0).

Thus, in a strongly balanced permutation, whenever the initial segment  $\{\pi(i)\}_{i=1}^k$  is ‘‘above average’’ with respect to a particular coordinate  $j$ , it can be made ‘‘below average’’ by flipping at most some fixed number  $l$  of items, and *vice versa*. As the name suggests, the strong balance condition is stricter than the usual balance condition. Nonetheless, the following lemma says that strongly balanced permutations always exist.

**Lemma 5.5** For any sequence  $\{w_i\}_{i=1}^m$  in  $\mathbf{R}^d$ , there exists a strongly  $16d^2$ -balanced permutation.

Note that this lemma claims only that a *single* strongly balanced permutation exists; unlike Theorem 5.3, it makes no claims regarding almost uniformity. The proof of the lemma relies heavily on the result of Grinberg and Sevast'yanov quoted earlier (Lemma 5.2); the proof is straightforward but rather technical, so we defer it to section 6.

### 5.3 A good flow

Now that we have multi-dimensional balanced almost uniform permutations and strongly balanced permutations, we are ready to construct a good flow.

**Lemma 5.6** Fix any number of knapsack constraints  $d$ . For arbitrary item weights, it is possible to construct a multicommodity flow  $f$  in  $G_\Omega$  with  $\mathcal{C}(f)$  bounded by a polynomial in  $n$  and  $\mathcal{L}(f) = O(n)$ .

**Proof:** Recall that we identify each vertex  $\mathbf{x} \in \Omega$  with the index set  $X = \{i : x_i = 1\}$ . Let  $\widehat{\Omega} = \{X \in \Omega : a^j(X) \leq b^j - 3c_d \max_{i \in X} a_i^j\}$ , where  $c_d$  is the constant in the construction of balanced almost uniform permutations as in Theorem 5.3. Our main goal will be to construct a flow  $\widehat{f}$  which, simultaneously for every  $X, Y \in \widehat{\Omega}$ , sends one unit of flow from  $X$  to  $Y$ . This flow will satisfy  $\mathcal{C}(\widehat{f}) \leq \text{poly}(n)$  and  $\mathcal{L}(\widehat{f}) = O(n)$ .

Note that, from any vertex  $X \in \Omega$ , we can obtain a vertex  $\widehat{X} \in \widehat{\Omega}$  by removing at most  $3dc_d$  items. Thus, we can use an approach similar to that in section 4.1 to extend  $\widehat{f}$  to a multicommodity flow  $f$  on the whole of  $\Omega$ , and  $f$  will satisfy  $\mathcal{C}(f) \leq O(n^{6dc_d}) \text{poly}(n) \leq \text{poly}'(n)$  and  $\mathcal{L}(f) \leq \mathcal{L}(\widehat{f}) + 6dc_d = O(n)$ .

It remains to define the flow  $\widehat{f}$  and show that it has the properties claimed. Fix  $X, Y \in \widehat{\Omega}$ . As in the one-dimensional case, the high-level idea is to route the flow from  $X$  to  $Y$  according to a balanced almost uniform permutation on the symmetric difference  $X \oplus Y$ . Recall from the one-dimensional case that we want both path endpoints to be far enough below the hyperplane before we apply the balanced permutation; recall also that we cannot in general ensure this by removing only a fixed number of items from  $X$  and  $Y$  (e.g., if the items in  $X$  are much larger than those in  $Y$ ). In the one-dimensional case (see section 4.2), we overcame this obstacle using the pre-processing random walk. Here we will use a different mechanism based on considering the “large” and the “small” items in  $X \oplus Y$  separately. This mechanism is cruder but more robust, and works in the multidimensional setting.

Let  $M = (M^1, \dots, M^d)$ , where  $M^j = \min\{\max_{i \in X} a_i^j, \max_{i \in Y} a_i^j\}$ . Let  $L = \{i \in X \oplus Y : a_i^j > M^j \text{ for some } j\}$  and  $S = (X \oplus Y) - L$ . ( $L$  and  $S$  are the “large” and “small” items respectively.) Let  $\{w_i\}_{i \in X \oplus Y}$  be an enumeration of the weights of the items in  $X \oplus Y$ , where weights from  $Y$  appear with a positive sign and weights from  $X$  appear with a negative sign. Let  $\mu_1 = \frac{1}{|L|} \sum_{i \in L} w_i$ , and let  $\mu_2 = \frac{1}{|S|} \sum_{i \in S} w_i$ . Let  $\pi_1$  be a permutation on  $L$  which is strongly  $16d^2$ -balanced with respect to the weights  $\{w_i\}_{i \in L}$ , and let  $\pi_2$  be a  $p_d(|S|)$ -uniform permutation which is  $c_d$ -balanced with respect to the weights  $\{w_i - \mu_2\}_{i \in S}$ . The existence of  $\pi_1$  and  $\pi_2$  is guaranteed by Lemma 5.5 and Theorem 5.3 respectively. To obtain  $\pi$ , we will interleave the strongly balanced permutation  $\pi_1$  and the balanced permutation  $\pi_2$ . The rule for interleaving will be as follows. Suppose that  $\pi(1), \dots, \pi(k)$  have already been assigned, and that  $\pi\{1, \dots, k\} = \pi_1\{1, \dots, k_1\} \cup \pi_2\{1, \dots, k_2\}$ . Now, either  $\frac{k_1}{k} \leq \frac{|L|}{|L|+|S|}$  or  $\frac{k_2}{k} < \frac{|S|}{|L|+|S|}$ , so we can define  $\pi(k+1)$  by

$$\pi(k+1) = \begin{cases} \pi_1(k_1+1), & \text{if } \frac{k_1}{k} \leq \frac{|L|}{|L|+|S|}; \\ \pi_2(k_2+1), & \text{if } \frac{k_2}{k} < \frac{|S|}{|L|+|S|}. \end{cases}$$

Now let  $\mu = \frac{1}{|X \oplus Y|} \sum_{i \in X \oplus Y} w_i = \frac{|L|\mu_1 + |S|\mu_2}{|L|+|S|}$ . We claim that  $\pi$  satisfies the following condition. Fix  $j$  and  $k$ . Then there exist sets of indices  $V_1$  and  $V_2$ , with  $|V_i \cap \{1, \dots, k\}| \leq 17d^2$ , such that

$$\sum_{i \in V_1} w_{\pi(i)}^j \leq (k-1)\mu^j + 3c_d M^j; \quad (26)$$

$$\sum_{i \in V_2} w_{\pi(i)}^j \geq (k-1)\mu^j - 3c_d M^j. \quad (27)$$

We will prove this in the case  $\mu_1^j \geq \mu_2^j$ ; if  $\mu_1^j < \mu_2^j$  the proof is similar. Again, let  $k_1 = |L \cap \pi\{1, \dots, k\}|$  and  $k_2 = |S \cap \pi\{1, \dots, k\}|$ , so that  $\pi\{1, \dots, k\} = \pi_1\{1, \dots, k_1\} \cup \pi_2\{1, \dots, k_2\}$ . The method of interleaving ensures that

$$\frac{k_1-1}{k-1} \leq \frac{|L|}{|L|+|S|}; \quad \frac{k_2-1}{k-1} \leq \frac{|S|}{|L|+|S|}.$$

Therefore, since  $\mu_1^j \geq \mu_2^j$ , we have

$$(k_1 - 1)\mu_1^j + k_2\mu_2^j \leq (k - 1)\mu^j; \quad (28)$$

$$k_1\mu_1^j + (k_2 - 1)\mu_2^j \geq (k - 1)\mu^j. \quad (29)$$

Clearly, the strong balance condition on  $\pi_1$  implies that there exist  $A, A'$ , with  $|A \oplus \{1, \dots, k_1\}| \leq 16d^2 + 1$  and similarly for  $A'$ , such that

$$\sum_{i \in A} w_{\pi_1(i)}^j \leq (k_1 - 1)\mu_1^j; \quad (30)$$

$$\sum_{i \in A'} w_{\pi_1(i)}^j \geq k_1\mu_1^j. \quad (31)$$

Also, by the balance condition on  $\pi_2$  we have

$$\sum_{i=1}^{k_2} w_{\pi_2(i)}^j \leq k_2\mu_2^j + c_d \max_{i \in S} \{|w_i^j - \mu_2^j|\} \leq k_2\mu_2^j + 3c_d M^j; \quad (32)$$

$$\sum_{i=1}^{k_2} w_{\pi_2(i)}^j \geq k_2\mu_2^j - c_d \max_{i \in S} \{|w_i^j - \mu_2^j|\} \geq (k_2 - 1)\mu_2^j - 3c_d M^j. \quad (33)$$

Now, let  $B = \pi^{-1}(\pi_1(A) \cup \pi_2\{1, \dots, k_2\})$  and  $B' = \pi^{-1}(\pi_1(A') \cup \pi_2\{1, \dots, k_2\})$ . Then we have  $|B \oplus \{1, \dots, k\}| \leq 16d^2 + 1 \leq 17d^2$ , and

$$\sum_{i \in B} w_{\pi(i)}^j = \sum_{i \in A} w_{\pi_1(i)}^j + \sum_{i=1}^{k_2} w_{\pi_2(i)}^j.$$

Exactly analogous relations hold with  $B, A$  replaced by  $B', A'$ . Combining this with equations (28)–(33) gives (26) and (27).

Now,  $\pi$  determines a path  $\{Z_i\}_{i=0}^{|X \oplus Y|}$  from  $X$  to  $Y$ , where  $Z_0 = X$  and  $Z_i = X \oplus \{\pi(1), \dots, \pi(i)\}$  for  $1 \leq i \leq |X \oplus Y|$ . This path might not stay in  $\Omega$ , but we can alter it slightly so that it does. Equations (26) and (27) imply that for every  $k$  and  $j$ , there exists a set of indices  $W_k^j$  with  $|W_k^j| \leq 34d^2$  such that

$$a^j(Z_k - W_k^j) \leq \max\{a^j(X), a^j(Y)\} + 3c_d M^j \leq b^j; \quad (34)$$

$$a^j(Z_k \cup W_k^j) \geq \min\{a^j(X), a^j(Y)\} - 3c_d M^j. \quad (35)$$

Let  $W_0 = \emptyset$  and for  $k = 1, \dots, |X \oplus Y|$ , let  $W_k = \cup_{j=1}^d W_k^j$ . Then, for all  $k$ ,  $|W_k| \leq 34d^3$ , and  $a(Z_k - W_k) \leq \mathbf{b}$ . For  $0 \leq k \leq |X \oplus Y|$ , define

$$\bar{Z}_k = Z_k - W_k.$$

Then each  $\bar{Z}_k \in \Omega$ . Our flow from  $X$  to  $Y$  will pass through each of the  $\bar{Z}_k$  in turn. To get from  $\bar{Z}_k$  to  $\bar{Z}_{k+1}$ , we perform the following steps:

1. Remove each item in  $\bar{Z}_k - (\bar{Z}_k \cap \bar{Z}_{k+1})$  in index order.
2. Add each item in  $\bar{Z}_{k+1} - (\bar{Z}_k \cap \bar{Z}_{k+1})$  in index order.

Define  $W_x = (W_k \cup W_{k+1}) \cap X$ , and  $W_y = (W_k \cup W_{k+1}) \cap Y$ . By analogy with section 4, for each intermediate point  $Z$  along the path define the ‘‘encoding’’  $\tilde{Z}$  by

$$\tilde{Z} = (X \oplus Y - Z) \cup (X \cap Y) - (W_k \cup W_{k+1}),$$

and let  $U = \pi\{1, \dots, k\}$ . In similar fashion to our earlier analysis one can see that, for a given  $Z$ ,  $X$  and  $Y$  are completely specified by the 4-tuple  $(\tilde{Z}, U, W_x, W_y)$ . We also have

$$\begin{aligned} a^j(\tilde{Z}) &= a^j(X) + a^j(Y) - a^j(Z \cup W_k \cup W_{k+1}) \\ &\leq a^j(X) + a^j(Y) - \min\{a^j(Z_k \cup W_k), a^j(Z_{k+1} \cup W_{k+1})\} \\ &\leq a^j(X) + a^j(Y) - (\min\{a^j(X), a^j(Y)\} - 3c_d M^j) \\ &= \max\{a^j(X), a^j(Y)\} + 3c_d M^j \\ &\leq b^j, \end{aligned}$$

where the second inequality follows from (35). Hence  $\tilde{Z} \in \Omega$ . We can therefore bound the flow  $\hat{f}(Z)$  through  $Z$  by

$$\hat{f}(Z) \leq \sum_{\tilde{Z} \in \Omega} \sum_{W_x, W_y, U} \Pr[\{\pi(1), \dots, \pi(|U|)\} = U]. \quad (36)$$

Finally, for a given  $X$  and  $Y$ , let  $L^j = \{i \in X \cup Y : a_i^j > M^j\}$ , so that  $L = \cup_{j=1}^d L^j$ . Note that for every  $j$ ,  $L^j \cap Y$  is equal to either  $L^j$  or  $\emptyset$ . Thus, if we define  $k_2 = |U \cap S|$ , then for given values of  $M$  and  $k_2$ , there are at most  $2^d \binom{|S|}{k_2}$  possible values for  $U$  in the inner sum of equation (36). Therefore, we have

$$\begin{aligned} \hat{f}(Z) &\leq \sum_{\tilde{Z} \in \Omega} \sum_{M, W_x, W_y, k_2} \sum_{U: |U \cap S| = k_2} \Pr[\{\pi_2(1), \dots, \pi_2(k_2)\} = U \cap S] \\ &\leq \sum_{\tilde{Z} \in \Omega} \sum_{M, W_x, W_y, k_2} \sum_{U: |U \cap S| = k_2} p_d(|S|) \binom{|S|}{k_2}^{-1} \\ &\leq \sum_{\tilde{Z} \in \Omega} \sum_{M, W_x, W_y, k_2} 2^d p_d(|S|) \\ &\leq \sum_{\tilde{Z} \in \Omega} n^d \left[ \binom{n}{68d^3} 2^{68d^3} \right] n \times 2^d p_d(n) \\ &= \text{poly}(n) |\Omega|, \end{aligned}$$

where in the second line we have appealed to the almost uniformity of permutation  $\pi_2$ . This completes the proof.  $\square$

Given such a flow, we can appeal to Theorem 2.1 to derive the main result of this section.

**Theorem 5.7** *Fix any dimension  $d > 0$ , and let  $\Omega$  be the set of solutions to an arbitrary instance of the  $d$ -dimensional 0-1 knapsack problem. The mixing time of the random walk on  $G_\Omega$  is  $\text{poly}_d(n)$ .*

As in one dimension, this immediately implies the existence of an fpras for computing  $|\Omega|$  in this more general setting.

**Remark:** In the multidimensional case we have made no attempt to optimize the exponent in our polynomial bound on the mixing time. Indeed, tracing through the proof of Lemma 5.6 we see that the bound is of the form  $n^{\mathcal{O}((3d)!)}$ . This is much larger than the lower bound  $n^{\Omega(d)}$  mentioned earlier. It would be interesting to try to close this gap.  $\square$

## 6 Proofs of technical lemmas

In this final section we supply the proofs of some technical results that were omitted from the main text.

### 6.1 Lemmas for section 3.1

The following two lemmas were used in the proof of Lemma 3.3.

**Lemma 6.1** *Let  $\nu$  be a uniform random permutation in  $\mathcal{S}_m$ . Then*

$$\Pr[\nu \text{ is not } \alpha\text{-good}] \leq 2m \exp(-2\alpha^2).$$

**Proof:** We adopt the notation of the proof of Lemma 3.3. Let  $1 \leq k_1 \leq m_1$ . It suffices to show that

$$\Pr\left[\left|\sum_{i=1}^{k_1} w_{\nu_1(i)} - k_1 \mu_1\right| > \alpha(M-1)\sqrt{k_1}\right] \leq 2 \exp(-2\alpha^2),$$

for then the lemma follows from the union bound and from symmetry (which allows us to replace  $k_1$  by  $k_1^* = \min\{k_1, m_1 - k_1\}$ ). But this inequality is a direct consequence of Hoeffding's bound on deviations in sampling without replacement [8].  $\square$

**Lemma 6.2** *Let  $m_1, m_2, k_1, k_2$  be non-negative integers and  $m = m_1 + m_2, k = k_1 + k_2$ . Then*

$$\frac{\binom{m_1}{k_1} \binom{m_2}{k_2}}{\binom{m}{k}} \geq A m^{-1/2} \exp\left\{-\frac{l^2 + \frac{1}{2}|l|}{\gamma(1-\gamma)} \left(\frac{1}{m_1} + \frac{1}{m_2}\right)\right\},$$

where  $\gamma = \frac{k}{m}, l = \frac{m_1 k_2 - m_2 k_1}{m}$ , and  $A > 0$  is a universal constant.

**Proof:** Note that  $k = \gamma m, k_1 = \gamma m_1 - l$ , and  $k_2 = \gamma m_2 + l$ . By the symmetry of binomial coefficients, we may assume that  $l \geq 0$ . We shall prove the lemma by showing the two inequalities

$$\frac{\binom{m_1}{\gamma m_1} \binom{m_2}{\gamma m_2}}{\binom{m}{\gamma m}} \geq A_1 m^{-1/2} \tag{37}$$

and

$$\frac{\binom{m_1}{\gamma m_1 - l} \binom{m_2}{\gamma m_2 + l}}{\binom{m_1}{\gamma m_1} \binom{m_2}{\gamma m_2}} \geq A_2 \exp\left\{-\frac{l^2 + \frac{1}{2}|l|}{\gamma(1-\gamma)} \left(\frac{1}{m_1} + \frac{1}{m_2}\right)\right\} \tag{38}$$

for positive constants  $A_1, A_2$ .

The first inequality is an immediate consequence of Stirling's approximation,  $\sqrt{2\pi n} \left(\frac{n}{e}\right)^n \leq n! \leq A_3 \sqrt{2\pi n} \left(\frac{n}{e}\right)^n$ , where  $A_3 = 1 + e^{1/12}$  is a constant. To prove the second inequality, we apply Stirling's approximation to all four binomial coefficients to get the following lower bound on the left-hand side of (38):

$$\left[ \frac{\mathrm{P}(\gamma m_1) \mathrm{P}((1-\gamma)m_1) \mathrm{P}(\gamma m_2) \mathrm{P}((1-\gamma)m_2)}{\mathrm{P}(\gamma m_1 - l) \mathrm{P}((1-\gamma)m_1 + l) \mathrm{P}(\gamma m_2 + l) \mathrm{P}((1-\gamma)m_2 - l)} \right] \times \left[ \frac{\gamma m_1 (1-\gamma)m_1 \gamma m_2 (1-\gamma)m_2}{(\gamma m_1 - l)((1-\gamma)m_1 + l)(\gamma m_2 + l)((1-\gamma)m_2 - l)} \right]^{1/2}, \tag{39}$$

where  $P(x)$  denotes  $x^x$ . Now we have

$$\begin{aligned} \frac{P(\gamma m_1)}{P(\gamma m_1 - l)} &= (\gamma m_1)^l \left(1 + \frac{l}{\gamma m_1 - l}\right)^{\gamma m_1 - l} \\ &\geq (\gamma m_1)^l \exp\left\{\frac{l(\gamma m_1 - l)}{\gamma m_1}\right\}, \end{aligned}$$

where we have used the inequality  $(1 + \frac{x}{y})^y \geq \exp(\frac{xy}{x+y})$ , valid for  $x, y > 0$ . Handling the three other pairs of factors in the numerator and denominator in similar fashion (using in addition the inequality  $(1 - \frac{x}{y})^y \geq \exp(\frac{-xy}{y-x})$ , valid for  $y > x > 0$ ) we see that the first parenthesis in (39) is bounded below by

$$\exp\left\{-\frac{l^2}{\gamma(1-\gamma)}\left(\frac{1}{m_1} + \frac{1}{m_2}\right)\right\}. \quad (40)$$

A similar calculation bounds the second parenthesis in (39) by

$$\exp\left\{-\frac{|l|}{\gamma(1-\gamma)}\left(\frac{1}{m_1} + \frac{1}{m_2}\right)\right\}. \quad (41)$$

Combining (40) and (41) completes the verification of inequality (38) above, and hence the proof of the lemma.  $\square$

## 6.2 Proof of Lemma 4.1

By removing  $X \cap Y$  from both  $X$  and  $Y$  and replacing  $b$  by  $b - a(X \cap Y)$ , we may assume that  $X \cap Y = \emptyset$ . Moreover, by scaling all the  $a_i$  and  $b$  we may assume that  $M = \max_i a_i = 1$ . Finally, we may assume that  $b \geq 3\Delta$  since otherwise there are no pairs  $(X, Y)$  satisfying the hypothesis of the lemma.

Define

$$\begin{aligned} \mathcal{F} &= \{(X', Y') : a(X') \geq b - \Delta \text{ or } a(Y') \geq b - \Delta\}; \\ \mathcal{E} &= \{(X', Y') : a(X') \leq b - 2\Delta \text{ and } a(Y') \leq b - 2\Delta\}. \end{aligned}$$

Note that  $\mathcal{F}$  contains all full pairs  $(X', Y')$ , and  $\mathcal{E}, \mathcal{F}$  are disjoint. Also, define the hitting times

$$\begin{aligned} \mathcal{T} &= \max_{(X', Y') \in \mathcal{F}} \text{E}(\text{number of PRW steps to hit } \mathcal{E} \text{ starting at } (X', Y')); \\ \mathcal{U} &= \min_{(X', Y') \in \mathcal{E}} \text{E}(\text{number of PRW steps to hit } \mathcal{F} \text{ starting at } (X', Y')). \end{aligned}$$

Now we claim that the lemma will follow if we can show:

- (i)  $\mathcal{T} \leq \alpha m$  for some constant  $\alpha > 0$ ;
- (ii)  $\mathcal{U}/\mathcal{T} \geq \beta$  for some constant  $\beta > 0$ .

To see this, set the length of the PRW to be  $C_1 m = 4\alpha m$ , and let  $(X_t, Y_t)$  denote the sequence of pairs visited by the PRW, with  $(X_0, Y_0) = (X, Y) \in \mathcal{F}$ . Let  $T_0$  be the first time  $t$  at which  $(X_t, Y_t) \in \mathcal{E}$  (or  $T_0 = C_1 m$  if the walk ends before this occurs); then let  $U_1$  be the first  $t$  for which  $(X_{T_0+t}, Y_{T_0+t}) \in \mathcal{F}$ , and  $T_1$  the first  $t$  for which  $(X_{T_0+U_1+t}, Y_{T_0+U_1+t}) \in \mathcal{E}$ . Continue defining a sequence of hitting times  $U_2, T_2, U_3, T_3, \dots$  in this way until the end of the walk is reached. Note

that the PRW is not full for at least  $\sum_i U_i$  steps, and that  $\sum_{i \geq 0} T_i + \sum_{i \geq 1} U_i = 4\alpha m$  is the total walk length. Now from facts (i) and (ii) we have

$$\mathbb{E}\left(\sum_{i \geq 0} T_i - \frac{1}{\beta} \sum_{i \geq 1} U_i\right) = \mathbb{E}\left(T_0 + \sum_{i \geq 1} (T_i - \frac{1}{\beta} U_i)\right) \leq \alpha m.$$

An application of Markov's inequality then ensures that  $\sum_{i \geq 0} T_i - \frac{1}{\beta} \sum_{i \geq 1} U_i \leq 2\alpha m$  with probability at least  $\frac{1}{2}$ . Conditioning on this event we have  $(1 + \frac{1}{\beta}) \sum U_i \geq 2\alpha m$ , and thus the proportion of steps during which the PRW is not full is at least  $1/2(1 + \frac{1}{\beta})$ , a constant. The lemma now follows easily.

It remains to verify facts (i) and (ii) above: these are immediate consequences of the following two claims. Let  $\sigma^2 = \frac{1}{m} \sum_{i \in X \cup Y} a_i^2$  be the second moment of the item weights, and note that  $\sigma^2 \geq 1/m$  since  $\max_i a_i = 1$ .

**Claim 1:**  $\mathcal{T} \leq \gamma_1/\sigma^2$  for a constant  $\gamma_1 > 0$ .

**Claim 2:**  $\mathcal{U} \geq \gamma_2/\sigma^2$  for a constant  $\gamma_2 > 0$ .

**Proof of Claim 1:** Choose an initial pair  $(X_0, Y_0) \in \mathcal{F}$  that maximizes the expected time until the PRW hits  $\mathcal{E}$ , and let  $(X_t, Y_t)$  denote the PRW starting at  $(X_0, Y_0)$ . We may assume w.l.o.g. that  $a(X_0) > a(Y_0)$ , so that  $a(X_0) \in [b - \Delta, b]$  and  $\mathcal{T}$  is the expected time until  $a(X_t) \leq b - 2\Delta$ . We estimate  $\mathcal{T}$  by coupling the PRW with the *unconstrained* random walk, which behaves exactly like the PRW except that the constraint  $\sum a_i \leq b$  is ignored. (Thus it is just simple random walk on an  $m$ -dimensional hypercube with holding probability  $\frac{1}{2}$  at every step.) Write  $(\hat{X}_t, \hat{Y}_t)$  for the unconstrained random walk, with  $(\hat{X}_0, \hat{Y}_0) = (X_0, Y_0)$ , and consider the first time  $t = \hat{T}$  at which  $|a(\hat{X}_t) - a(X_0)| \geq 2\Delta$ . Now  $a(\hat{X}_t)$  is a supermartingale up to time  $\hat{T}$ , since  $\mathbb{E}(a(\hat{X}_{t+1}) - a(\hat{X}_t) | \hat{X}_t) = \frac{1}{2m}(a(\hat{Y}_t) - a(\hat{X}_t)) < 0$ . Thus with constant probability  $a(\hat{X}_{\hat{T}}) \leq a(\hat{X}_0) - 2\Delta$ , and so  $(X_{\hat{T}}, Y_{\hat{T}}) \in \mathcal{E}$ . Hence  $\mathcal{T}$  is bounded above by a constant times  $\mathbb{E}(\hat{T})$ . But we also have  $\mathbb{E}((a(\hat{X}_{t+1}) - a(\hat{X}_t))^2 | \hat{X}_t) = \frac{1}{2m} \sum_i a_i^2 = \frac{\sigma^2}{2}$ . So  $\mathbb{E}(\hat{T})$  is the expected time for a supermartingale with increments bounded by  $\pm 1$  and with second moment  $\sigma^2/2$  to move a distance  $\pm 2\Delta$  from its initial value. A standard application of the martingale Optional Stopping Theorem (see, e.g., [6, Section 12.5]), now yields that  $\mathbb{E}(\hat{T}) \leq \frac{(4\Delta+1)^2}{\sigma^2/2} = \gamma_1/\sigma^2$  for a positive constant  $\gamma_1$ . This completes the proof of Claim 1.

**Proof of Claim 2:** As above let  $(X_t, Y_t)$  denote the PRW, but now with  $(X_0, Y_0) \in \mathcal{E}$ . We follow the random variable  $Z_t = a(X_t) - a(Y_t)$ , which always has a drift towards 0 (i.e.,  $\mathbb{E}(Z_{t+1} - Z_t | X_t) \times Z_t \leq 0$  for all  $t$ ). Note that initially  $|Z_0| \leq 2(b - 2\Delta) - V$ , where  $V = a(X_t) + a(Y_t) = \sum_i a_i$  is independent of  $t$ . And when  $(X_t, Y_t) \in \mathcal{F}$  we have  $|Z_t| \geq 2(b - \Delta) - V$ . Thus  $\mathcal{U}$  is bounded below by the minimum expected time for  $|Z_t|$  to increase by  $2\Delta$  from its initial value. But the second moment is  $\mathbb{E}((Z_{t+1} - Z_t)^2 | X_t) = \frac{1}{2m} \sum_i (2a_i)^2 = 2\sigma^2$ , so by a similar application of the Optional Stopping Theorem we conclude that  $\mathcal{U} \geq \gamma_2/\sigma^2$ , as claimed.

This completes the verification of Claims 1 and 2, and hence the proof of the lemma.  $\square$

### 6.3 Proof of Lemma 5.5

First suppose that  $\sum_{i=1}^m w_i = 0$ . We will show that, in this case, there exists a strongly  $8d^2$ -balanced permutation  $\pi$ . Let  $L$  be the set containing the  $4d$  indices  $i$  with the largest values of  $w_i^j$ , and the  $4d$  indices  $i$  with the largest values of  $-w_i^j$ , for each  $j \leq d$ . Then  $|L| \leq 8d^2$ .

The permutation  $\pi$  we construct will satisfy  $\{\pi(m), \dots, \pi(m - |L| + 1)\} = L$ . It will be enough to check that the strong balance condition holds for  $1 \leq k \leq m - |L|$ . It suffices to show that, for

all  $j \leq d$  and  $k \leq m - |L|$ , we have

$$-s^{j+} \leq \sum_{i=1}^k w_{\pi(i)}^j \leq s^{j-}, \quad (42)$$

where

$$s^{j+} \equiv \sum_{i \in L} (w_i^j)^+; \quad s^{j-} \equiv \sum_{i \in L} (w_i^j)^-.$$

We will need the Grinberg-Sevast'yanov result (Lemma 5.2), which states that for any set of vectors  $x_1, \dots, x_n$  in  $\mathbf{R}^d$  with  $\sum_i x_i = 0$ , there exists a permutation  $\nu \in \mathcal{S}_n$  such that

$$\sum_{i=1}^k x_{\nu(i)} \in d \times \text{conv}\{x_1, \dots, x_n\} \quad \text{for } 1 \leq k \leq n.$$

Note that the permutation  $\nu'$  defined by  $\nu'(i) = \nu(n+1-i)$  for all  $i$  satisfies

$$\sum_{i=1}^k x_{\nu'(i)} \in -d \times \text{conv}\{x_1, \dots, x_n\} \quad \text{for } 1 \leq k \leq n,$$

because  $\sum_{i=1}^k x_{\nu'(i)} = \sum_{i=m-k+1}^m x_{\nu(i)} = -\sum_{i=1}^k x_{\nu(i)}$ .

Let  $S_1 = \{1, \dots, m\} - L$ , let  $m' = m - |L|$ , and let  $\pi_1$  be a permutation on  $S_1$  such that

$$\sum_{i=1}^k (w_{\pi_1(i)} - \mu_1) \in -d \times \text{conv}\{w_i - \mu_1 : i \in S_1\},$$

for all  $k$ , where  $\mu_1 = \frac{1}{m'} \sum_{i \in S_1} w_i$ . Suppose that  $m'$  is even and  $m' = 2r$ ; if  $m'$  is odd the proof is similar. Now, let  $S_2 = \{\pi_1(r+1), \dots, \pi_1(m')\}$ , and let  $\pi_2$  be a permutation on  $S_2$  such that

$$\sum_{i=1}^k (w_{\pi_2(i)} - \mu_2) \in d \times \text{conv}\{w_i - \mu_2 : i \in S_2\}, \quad (43)$$

where  $\mu_2 = \frac{1}{r} \sum_{i \in S_2} w_i$ . Define the permutation  $\pi$  by

$$\pi(i) = \begin{cases} \pi_1(i), & \text{if } i \leq r; \\ \pi_2(i-r), & \text{if } r < i \leq m'. \end{cases}$$

We must check that  $\pi$  satisfies (42). Fix  $j$ . W.l.o.g.  $s^{j+} \leq s^{j-}$ , so that  $\mu_1^j \geq 0$ . For  $k \leq r$  we have

$$\begin{aligned} \sum_{i=1}^k w_{\pi(i)}^j &= \sum_{i=1}^k w_{\pi_1(i)}^j \\ &\geq \sum_{i=1}^k (w_{\pi_1(i)}^j - \mu_1^j) \\ &\geq -d \max_{1 \leq i \leq m'} \{w_i^j - \mu_1^j\} \\ &\geq -d \max_{1 \leq i \leq m'} \{w_i^j\} \\ &\geq -s^{j+}/4 \\ &\geq -s^{j+}, \end{aligned}$$

and

$$\begin{aligned}
\sum_{i=1}^k w_{\pi(i)}^j &= \sum_{i=1}^k w_{\pi_1(i)}^j \\
&= k\mu_1^j + \sum_{i=1}^k (w_{\pi_1(i)} - \mu_1^j) \\
&\leq r\mu_1^j + d \max_{1 \leq i \leq m'} \{-(w_i^j - \mu_1^j)\} \\
&= \frac{1}{2}(s^{j-} - s^{j+}) + d \max_{1 \leq i \leq m'} \{-w_i^j\} + d\mu_1^j \\
&\leq \frac{1}{2}(s^{j-} - s^{j+}) + (\frac{1}{4}s^{j-}) + (\frac{1}{4}s^{j+}) \\
&= \frac{3}{4}s^{j-} - \frac{1}{4}s^{j+} \\
&\leq s^{j-} - \frac{1}{4}s^{j+}.
\end{aligned} \tag{44}$$

(We will need the extra  $-\frac{1}{4}s^{j+}$  in the second part of the proof.) For  $r < k \leq 2r$  we have

$$\sum_{i=1}^k w_{\pi(i)}^j = \sum_{i=1}^r w_{\pi_1(i)}^j + \sum_{i=1}^{k-r} w_{\pi_2(i)}^j. \tag{45}$$

Now if  $\mu_2^j < 0$  then the conditional expectation of  $\sum_{i=1}^k w_{\pi(i)}^j$  given  $\pi_1$  is at least  $-s^{j+} + s^{j-}$ . Hence we must have

$$\begin{aligned}
\sum_{i=1}^k w_{\pi(i)}^j &\geq -s^{j+} + s^{j-} - d \max_{i \in S_2} \{-(w_i^j - \mu_2^j)\} \\
&\geq -s^{j+} + s^{j-} - d \max_{i \in S_2} \{-w_i^j\} \\
&\geq -s^{j+} + s^{j-} - s^{j-}/4 \\
&\geq -s^{j+}.
\end{aligned}$$

On the other hand, if  $\mu_2^j \geq 0$  the right-hand side of (45) can be bounded below as follows:

$$\begin{aligned}
\sum_{i=1}^r w_{\pi_1(i)}^j + \sum_{i=1}^{k-r} w_{\pi_2(i)}^j &= \left[ \sum_{i=1}^r w_{\pi_1(i)}^j + (k-r)\mu_2^j \right] + \sum_{i=1}^{k-r} (w_{\pi_2(i)}^j - \mu_2^j) \\
&\geq \left[ \frac{1}{2}(s^{j-} - s^{j+}) - d \max_{1 \leq i \leq m'} \{w_i^j - \mu_1^j\} \right] - d \max_{i \in S_2} \{-(w_i^j - \mu_2^j)\} \\
&\geq \left[ \frac{1}{2}(s^{j-} - s^{j+}) - d \max_{1 \leq i \leq m'} \{w_i^j\} \right] - d \max_{i \in S_2} \{-w_i^j\} - d\mu_2^j \\
&\geq \left[ \frac{1}{2}(s^{j-} - s^{j+}) - s^{j+}/4 \right] - s^{j-}/4 - s^{j-}/4 \\
&= -\frac{3}{4}s^{j+} \\
&\geq -s^{j+}.
\end{aligned}$$

For a corresponding upper bound, we can write

$$\sum_{i=1}^k w_{\pi(i)}^j \leq \sum_{i=1}^r w_{\pi_1(i)}^j + (k-r)\mu_2^j + \min\{d, k-r\} \max_{i \in S_2} \{w_i^j - \mu_2^j\}. \tag{46}$$

If  $\mu_2^j \geq 0$ , the right-hand side of (46) is bounded above by

$$(s^{j^-} - s^{j^+}) + d \max_{i \in S_2} \{w_i^j\} \leq (s^{j^-} - s^{j^+}) + s^{j^+}/4 \leq s^{j^-}.$$

If  $\mu_2^j < 0$ , the right-hand side of (46) is bounded above by

$$\sum_{i=1}^r w_{\pi_1(i)}^j + \min\{d, k - r\} \max_{i \in S_2} \{w_i^j\} \leq (s^{j^-} - s^{j^+}/4) + s^{j^+}/4 \leq s^{j^-},$$

where in the first inequality we have used equation (44).

Putting all the above together we see that  $\pi$  is strongly  $8d^2$ -balanced. Furthermore, in light of (42) we know also that, for all  $k$  and  $j$ , there exists a set of indices  $S \supseteq \{1, \dots, k\}$ , with  $|S| \leq k + 8d^2$ , such that  $\sum_{i=1}^k w_{\pi(i)}^j$  and  $\sum_{i \in S} w_{\pi(i)}^j$  have opposite signs.

Now let  $\{w_i\}_{i=1}^m$  be arbitrary. From the above, there exists a permutation  $\pi$  which is strongly  $8d^2$ -balanced with respect to the sequence  $\{w_i - \mu\}_{i=1}^m$ , where  $\mu = \frac{1}{m} \sum_{i=1}^m w_i$ . We claim that  $\pi$  is also strongly  $16d^2$ -balanced with respect to the original sequence  $\{w_i\}_{i=1}^m$ . Fix  $k$  and  $j$ , and define  $\mu^j = \frac{1}{m} \sum_{i=1}^m w_i^j$ . W.l.o.g.  $\mu^j \geq 0$ . Then there exists  $S \supseteq \{1, \dots, k\}$  with  $|S| \leq k + 8d^2$  such that

$$\sum_{i \in S} w_{\pi(i)}^j - k\mu^j \geq \sum_{i \in S} w_{\pi(i)}^j - |S|\mu^j \geq 0.$$

In addition, there exists  $S \supseteq \{1, \dots, k\}$ , with  $|S| \leq k + 8d^2$ , such that  $\sum_{i \in S} w_{\pi(i)}^j \leq |S|\mu^j$ . It follows that, for some  $S' \subseteq S$  with  $|S'| = k$ , we have  $\sum_{i \in S'} w_{\pi(i)}^j \leq k\mu^j$ . Since  $|S' \oplus \{1, \dots, k\}| \leq 16d^2$  this completes the proof.  $\square$

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