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2-SAT: can solve in poly-time via a reduction to strongly connected components of directed graphs.

Here is a simple randomized algorithm.

1. Start with arbitrary assignment.
2. Repeat $\leq 100 n^2$ times: (unless all clauses are satisfied & then stop)
 - choose an arbitrary unsatisfied clause C
 - Pick a literal in C v.a.r. & satisfy that literal.
3. If we are at a satisfying assignment, output it. Else, output "unsatisfiable."

Fix a satisfying assignment, call it τ .

Let σ_t be the assignment of the alg. at round t .

Let $X_t = \#$ of variables that agree
between σ_t & τ .

If $X_t = n$ then the alg. found a satisfying
assignment.

$X_t \in \{0, 1, \dots, n\}$ & is a random walk.

Claim: $\Pr(X_{t+1} = i+1 \mid X_t = i) \geq \frac{1}{2}$.

Why?

Consider the unsatisfied clause C updated $X_t \rightarrow X_{t+1}$

We know τ satisfies C , hence

≥ 1 of the 2 variables in C

have opposite assignment in τ vs. σ_t

& ≤ 1 of the 2 agree in σ_t & τ .

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Consider slowed down version (Y_t) where:

$$Y_0 = X_0 \quad \& \quad \Pr(Y_{t+1} = i+1 \mid Y_t = i) = \frac{1}{2}$$

"Couple" the 2 processes (X_t) & (Y_t)

So that if $Y_{t+1} = Y_t + 1$ then $X_{t+1} = X_t + 1$
or $X_t = n$.

More precisely, if $Y_t \rightarrow Y_{t+1}$ increases

then $X_t \rightarrow X_{t+1}$ chooses ≥ 1 of the
"good" variables in the chosen clause
to update.

Y_t is an unbiased random walk on $\{0, 1, \dots, n\}$.

Claim: Expected time to hit = reach n is $O(n^2)$.

Let $h_j =$ expected # of steps to reach n
starting at $X_0 = j$.

For $0 \leq j < n$,

Note,
$$h_j = \frac{1}{2}h_{j-1} + \frac{1}{2}h_{j+1} + 1$$

hence,
$$h_j - h_{j+1} = h_{j-1} - h_j + 2$$

$$h_n = 0$$

$$h_0 - h_1 = 1$$

By induction,
$$h_j - h_{j+1} = 2j + 1$$

Therefore,

$$\begin{aligned} h_0 - h_n &= \sum_{i=0}^{n-1} (h_i - h_{i+1}) \\ &= \sum_{i=0}^{n-1} (2i + 1) = \frac{2n(n-1)}{2} + n = n^2. \end{aligned}$$

Finite Markov chains:

State space \mathcal{S}

consider $\mathcal{S} = \{0, 1, \dots, N-1\}$. (often N is huge)

Think of graph on \mathcal{S}

directed edges s.t. for each $i \in \mathcal{S}$,

$$\sum_{j \in \mathcal{S}} P(i,j) = 1$$

hence, $P(i,j) = \Pr(X_{t+1}=j | X_t=i)$

P is $N \times N$ transition matrix

(needs to be stochastic = rows sum to 1)

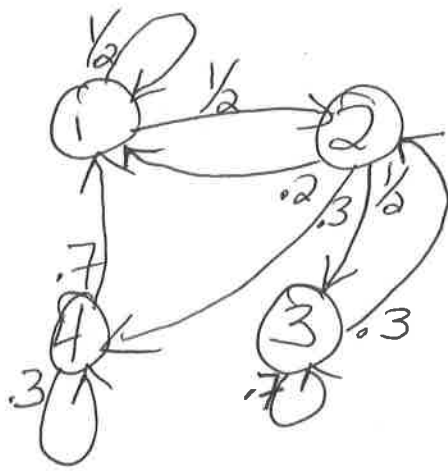
P^t is t -step probabilities:

$$\Pr(X_t=j | X_0=i) = P^t(i,j).$$

If $X_0 \sim \mu_0$ then $X_t \sim \mu_t$ where $\mu_t = \mu_0 P^t$

$\mu_0 = [\quad]$
row vector

Example:



$$P = \begin{bmatrix} .5 & .5 & 0 & 0 \\ .2 & 0 & .5 & .3 \\ 0 & .3 & .7 & 0 \\ .7 & 0 & 0 & .3 \end{bmatrix}$$

Note, $P^{20} = \begin{bmatrix} .244190 & .244187 & .406971 & .104652 \\ .244187 & .244186 & .406975 & .104651 \\ .244181 & .244185 & .406984 & .104650 \\ .244195 & .244188 & .406966 & .104652 \end{bmatrix}$

for distribution $\pi \approx [.2442, .2442, .4070, .10465]$

$$\lim_{t \rightarrow \infty} P^t = \begin{bmatrix} \pi \\ \pi \\ \pi \\ \pi \end{bmatrix}$$

A stationary distribution satisfies: $\pi = \pi P$
i.e., invariant wrt transition matrix,
like a fixed point.

(once you're in π , it stays in π)

π is an eigenvector of P with eigenvalue 1.

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Ergodic: if $\exists t$ s.t. $\forall i, j \in \mathcal{S}, P^t(i, j) > 0$
(graph defined by P^t is fully-connected)

Irreducible: if $\forall i, j \in \mathcal{S}, \exists t$ s.t. $P^t(i, j) > 0$
(graph defined by P is 1 scc)

for state $i \in \mathcal{S}$, period of $i = \gcd \{ t : P^t(i, i) > 0 \}$

Aperiodic: period of all $i \in \mathcal{S}$ is 1.

Ergodic \iff Irreducible & aperiodic.

Theorem: For a finite ergodic MC,
 there is a unique stationary distribution π
 & for all $i, j \in \Omega$,

$$\lim_{t \rightarrow \infty} P^t(i, j) = \pi(j).$$

(in words: no matter the initial distribution μ_0 ,
 $\lim_{t \rightarrow \infty} \mu_t = \pi$)

What is π ?

In general need to Gaussian elimination
 to find it but usually $|\Omega|$ is HUGE.

If P is symmetric then $\pi = \text{uniform}(\Omega)$.

Proof: Need to verify for $\pi(i) = \frac{1}{N}$ then $\pi P = \pi$.

$$\begin{aligned} (\pi P)(i) &= \sum_{k \in \Omega} \pi(k) P(k, i) \\ &= \frac{1}{N} \sum_k P(k, i) = \frac{1}{N} \sum_k P(i, k) = \frac{1}{N} \quad \square \end{aligned}$$

↑ since P is symmetric
↑ since P is stochastic

Weighted symmetric:

P is reversible with respect to π if:

$$\forall i, j \in \mathbb{Z}, \pi(i)P(i, j) = \pi(j)P(j, i)$$

Such a π is a stationary distribution.

Proof: $(\pi P)(i) = \sum_{k \in \mathbb{Z}} \pi(k)P(k, i) = \sum_k \pi(i)P(i, k) = \pi(i) \sum_k P(i, k) = \pi(i)$

Random walk on d -regular undirected graph:

for edge (i, j) , $P(i, j) = P(j, i) = \frac{1}{d}$

so it's symmetric & $\pi(i) = \frac{1}{n}$ for $n = |V|$.

Non-regular?

Then $\pi(i) = \frac{d(i)}{Z}$ where $d(i) = \text{degree of } i$
 $Z = \sum_j d(j) = 2m$

Check: $\pi(i)P(i, j) = \frac{d(i)}{Z} \frac{1}{d(i)} = \frac{1}{Z} = \pi(j)P(j, i)$.

What if G is directed? No idea about $\pi!$

Proof that ergodic, finite MC has a stationary distribution.

Can prove using Perron-Frobenius Theorem.

Constructive proof:

Let $h_{ij} = E[T_{ij}] =$ expected hitting time

where $T_{ij} = \min\{t : X_t = j \mid X_0 = i\}$

Lemma: $\pi(i) = \frac{1}{h_{ii}}$ where $h_{ii} =$ expected 1st return time for state i .

Claim: $h_{ij} < \infty$

Proof: Since P is ergodic, $\exists t^* \& \epsilon > 0$ s.t.

$$\forall k, l \in \mathcal{J}, P^{t^*}(k, l) \geq \epsilon.$$

Set $X_0 = i$.

$$Pr(X_{t^*} = j \mid X_0 = i) \geq \epsilon.$$

$$\& Pr(X_{2t^*} = j \mid X_{t^*}) \geq \epsilon.$$

Thus, $Pr(\forall t \leq t^* l, X_t \neq j) \leq (1-\epsilon)^l \leq e^{-\epsilon l} \rightarrow 0$

$$Pr(T_{ij} > t^* l) \geq 1 - (1-\epsilon)^l \geq 1 - e^{-\epsilon l} \rightarrow 1. \quad \square$$

PageRank:

Method to assign "importance" to webpages.

Graph where $V = \text{webpages}$

$E = \text{directed edges corresponding to hyperlinks.}$

Idea 1: a link is a citation, so count # of in-edges.

Idea 2: Weight outgoing links by # of hyperlinks on it.
So if page x has d outgoing links then each gets $\frac{1}{d}$ of a citation.
Hence, it is like a random walk.

$$\pi(y) = \sum_{x: \vec{xy} \in E} \frac{1}{d(x)}$$

Idea 3: weight a page by its $\pi(x)$, hence:

$$\pi(y) = \sum_{x: \vec{xy} \in E} \frac{\pi(x)}{d(x)}$$

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This corresponds to the stationary distribution of the random walk on the web graph.

But what is π ? is it unique? Not necessarily because it may not be ergodic.

How to make it ergodic?

Choose $\alpha < 1$,

From page $x \in V$,

with prob. α , choose a random out edge
with prob. $1-\alpha$, choose a random vertex
in whole graph.

Then clearly ergodic so unique π .

But what is π ?

This is the PageRank vector.

Metropolis filter:

for $x \in \mathcal{X}$, have weight $w(x) > 0$

want to design MC whose stationary distribution π

satisfies $\pi(x) \propto w(x)$

in other words, $\pi(x) = \frac{w(x)}{\sum_{y \in \mathcal{X}} w(y)} = Z$

Choose transitions so that the graph $\mathcal{G}(\mathcal{X}, P)$ is strongly connected, but what probabilities $P(i,j)$?

Consider transition $X_{old} \rightarrow X_{new}$:

set $P(X_{old}, X_{new}) = \min\left\{1, \frac{w(X_{new})}{w(X_{old})}\right\}$

check reversibility: assume $w(X_{old}) \leq w(X_{new})$

$\pi(X_{old})P(X_{old}, X_{new}) = \frac{w(X_{old})}{Z} \times 1 = \frac{w(X_{old})}{Z}$ (Since)

$\& \pi(X_{new})P(X_{new}, X_{old}) = \frac{w(X_{new})}{Z} \times \frac{w(X_{old})}{w(X_{new})} = \frac{w(X_{old})}{Z}$